# Socioeconomic Determinants of Infant Mortality Rate Disparities 

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# Socioeconomic Determinants of Infant Mortality Rate Disparities 

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\begin{array}{c}\text { A Thesis } \\
\text { Presented to } \\
\text { the Graduate School of } \\
\text { Clemson University }\end{array}
$$\right] \begin{array}{c}In Partial Fulfillment of the <br>
Requirements for the Degree of <br>
Master of Arts in <br>

Economics\end{array}\right]\)| By |
| :---: |
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Accepted by<br>Dr. Thomas Mroz, Committee Chair<br>Dr. Michael Maloney<br>Dr. Raymond Sauer


#### Abstract

This study examines variation in infant mortality rate (IMR) between black and white populations. Data were drawn from CDC Birth Cohort Linked Birth Infant Data for 2004. Author examined numerous socioeconomic factors that could explain the gap in IMR between two races. Each potential factor was examined separately before introducing into a complete model. The proposed model explains about 11 percentage points in difference between groups, and showed factors that can be affected by change in public policy, play statistically significant role in IMR disparity. We can look at the IMR as a socioeconomic indicator that highlights both negative and positive developments in society.


Keywords: infant mortality rate, racial disparities, socioeconomic factors

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## Introduction

"No indicator captures the divergence in human development opportunity more powerfully than child mortality" ${ }^{1}$

Infant mortality rate or IMR is the number of newborns dying under a year of age divided by the number of live births that year. The infant mortality rate is also called the infant death rate. It is the number of deaths that occur in the first year of life for 1000 live births ${ }^{2}$.

In the past, infant mortality claimed a considerable percentage of children born, but the rate declined significantly, mainly due to improvements in basic health care, advances in the medical field and the introduction of prenatal care. Infant mortality rate is commonly included as a part of standard of living evaluations in economics ${ }^{3}$

The IMR is a useful indicator of a country's level of health or development, and is a component of the physical quality of life index. But the method of calculating IMR often varies widely between countries based on the way they define a live birth and how many premature infants are born in the country. The World Health Organization (WHO) defines a live birth as any born human being who demonstrates independent signs of life, including breathing, voluntary muscle movement, or heartbeat. Many countries, however, including certain European states and Japan, only count as live births cases where an infant breathes at birth, which makes their reported IMR numbers somewhat lower and raises their rates of perinatal mortality. The exclusion of any high-risk infants from the denominator or numerator in reported IMRs can be problematic for comparisons. Many countries, including the United States, Sweden or Germany, count an infant exhibiting any sign of life as alive, no matter the month of gestation or the size, but according to United States Center for Disease Control researchers ${ }^{4}$, some other countries differ in these practices. All of the countries named adopted the WHO definitions in the late 1980s or early 1990s ${ }^{5}$, which are used throughout the European Union.

[^0]My research centered on two US major groups: African Americans and Caucasians ${ }^{6}$. I want to explain significant gap in infant mortality rate between two groups in terms of socioeconomic factors. Unlike health factors, socioeconomic factors can be affected through the appropriate public policy. ${ }^{7}$ My study shows that some factors have an influence on one group and almost insignificant for another. If policy maker could take into account those differences we could allocate resources more efficiently, reducing disparity between groups and achieving lower IMR in U.S.

## Literature review

In his study of IMR in large urban areas Sims et al. (2007) ${ }^{8}$ revealed that high poverty was significantly associated with minority-white IMR disparities. Findings from this study suggested that the factors associated with infant mortality in urban areas vary by race and ethnicity.

Almond et al. (2008) ${ }^{9}$ showed that rolling out food stamps program during 1960s early 1970s helped improved birth outcomes for both whites and African Americans, with larger impacts for births to African American mothers.

Singh and Kogan (2007) ${ }^{10}$ analyzed vital records data between 1969 and 2001, and revealed that relative socioeconomic disparities in infant mortality, increased substantially since 1985. Improvements in infant mortality would be substantial if infants in the lower socioeconomic status groups experienced mortality rates similar to those of the highest status group. The key risk factors include: smoking during pregnancy, delayed or no prenatal care, and lack of health care coverage.

[^1]Farley et al. (2006) ${ }^{11}$ in his work showed that after controlling for individual level socio-demographic factors, median household income was positively associated with both birthweight-for-gestational-age and gestational age at birth.

Sohler et al. (2003) looked for possible association between income inequality and IMR in New York City. He showed that an increase of one standard deviation in income inequality was associated with an increase of 0.8 in IMR. ${ }^{12}$ Investigating economic inequality among 10 richest nations between 1903 and 2003 Leigh and Jencks (2006) ${ }^{13}$ found that income share of the top $10 \%$ of population is negatively related to life expectancy and positively related to infant mortality.

Sims and Rainge (2002) ${ }^{14}$ in their study of white and black mothers in Milwaukee found that black infant neonatal mortality rates were twice those of whites, while postneonatal mortality rates were three times that of whites. All black mothers were nearly eight times as likely as all white mothers to have inadequate prenatal care, whereas poor black mothers were three times as likely to have inadequate prenatal care as were poor white mothers.

In 2009, the US Center for Disease Control and Prevention (CDC) issued a report which stated that the American rates of infant mortality were affected by the United States' high rates of premature babies compared to European countries and which highlights the differences in reporting requirements between the United States and Europe. France, the Czech Republic, Ireland, the Netherlands, and Poland do not report all live births of babies under 500 g and/or 22 weeks of gestation. However, the report also concludes that the differences in reporting are unlikely to be the primary explanation for the United States' relatively low international ranking. ${ }^{15}$

[^2]
## Datasets

In the paper I primarily focus on individual level data. I also use county level of data in an auxiliary role to get further insight and test a few hypotheses, where individual level does not provide relevant information. Statistical package Stata ${ }^{16}$ was used to analyze the dataset.

Individual Level Dataset is available for public use by Center for Disease Control and Prevention. For this study I use latest available for Birth Cohort Linked Birth/Infant Data for 2004 ${ }^{17}$. The dataset has $4,118,956$ records for all births occurred in 2004 in U.S. with 27,763 deaths of infants born in the same year ${ }^{18}$. There are $3,839,003$ records of mothers who reported their race as Black or White with corresponding 26,278 deaths. Unfortunately, dataset does not have poverty level, income or insurance coverage for individual records. I tried to pick-up some correlation of income from education level of the mother, and county poverty level.

County level datasets were taken from U.S Census Bureau's County and City Data Book ${ }^{19}$. I used datasets for 2004, 2005 to match periods with my individual level data. After extracting and tabulating county datasets I merged them with my primary, individual level, dataset.

To help me explore the data, I created list of my own variables. Most of variables created are dummy group variables. They serve to define and combine individual records into groups. Group categories include race, age, residence, level of education attained; trimester of prenatal care began, etc. Grouping helped to look at the data from perspective of public policy maker where groups could be treated as separate homogeneous entities. For example we could look only at black mothers with education level less than high-school from age 22 to 25 , who started prenatal care in the third trimester.

Another reason for grouping is that dataset uses two different revisions (1989 and 2003 Revisions) of the U.S. Standard Certificate of Live Birth and Certificate of Death. Seven states implemented the revised certificate by January 1, 2004 and two during 2004. Comparable revised data was combined with data from the remaining states and the $D C$.

[^3]Groups were defined in such way that created group variables did not distort underlying information and made groups fully compatible and inclusive for both the unrevised and the revised version of certificates. As time goes on more states will switch to revised version, but as long as some states continue to use the unrevised version, such grouping will provide relevant information for all states, with no need to do separate data analysis for 'revised' and 'unrevised' states.

In 2004 infant mortality rate for blacks and whites was 6.84 per 1000, with significant disparity between two groups. While whites IMR was 5.63, blacks IMR was 13.19. Average age of mother was 27.34 and was slightly higher for white mothers 27.67 compare to black mothers 25.58 . Infants from white mothers accounted for $84 \%$ of the sample and remainder $16 \%$ were black. In $69.5 \%$ cases white mothers reported of being married compare to only $31.2 \%$. In the sample $21.8 \%$ had less than high school education, $29.5 \%$ high school education, and $47.4 \%$ more than high school. $79.9 \%$ started prenatal care in first trimester ( $70.1 \%$ of blacks and 81.7 of whites), $13.5 \%$ started in second trimester ( $19.2 \%$ of blacks and $12.3 \%$ of whites) and $2.85 \%$ started in third trimester ( $4.4 \%$ of blacks and $2.5 \%$ of whites). $1.067 \%$ didn't use prenatal care ( $2.025 \%$ of blacks and $0.883 \%$ of whites). $10.87 \%$ smoked during the pregnancy (11.45\% of whites and $8.05 \%$ of blacks). Neonatal mortality rate was 4.55 per 1000 births and postneonatal mortality rate 2.29 per 1000 births. Summary statistics for variables I use in the paper listed below ${ }^{20}$.

Table 1.1 Summary Statistics

| Variable | Mean | Std. Dev. | Min | Max | Observations |
| :--- | :--- | :--- | :--- | :--- | :--- |
| Outcome per 1000 birth | 6.845006 | 82.45092 | 0 | 1000 |  |
| IMR per 1000 blacks | 13.18835 | 114.0808 | 0 | 1000 | 616074 |
| IMR per 1000 whites | 5.632454 | 74.83803 | 0 | 1000 | 3222929 |
| Mother's age | 27.3396 | 6.18754 | 12 | 50 |  |
| Mother age, black | 25.58799 | 6.235583 | 12 | 50 | 616074 |
| Mother age, white | 27.67443 | 6.12 | 12 | 50 | 3222929 |
| Black Mother | .1604776 | .3670485 | 0 | 1 |  |
| White Mother | .8395224 | .3670485 | 0 | 1 |  |
| Married | .6333918 | .4818783 | 0 | 1 |  |
| Married blacks | .3118521 | .4632502 | 0 | 1 | 616074 |
| Married whites | .6948552 | .4604688 | 0 | 1 | 3222929 |
| Age 11 to 16 | .0167723 | .1284174 | 0 | 1 |  |
| Age 17 to 18 | .047957 | .2136753 | 0 | 1 |  |
| Age 19 to 21 | .1373148 | .3441794 | 0 | 1 |  |

${ }^{20}$ State residence dummy variables are not listed

Table 1.1 Summary Statistics (continue)

| Variable | Mean | Std. Dev. | Min | Max | Observations |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Age 22 to 25 | . 2154617 | . 4111423 | 0 | 1 |  |
| Age 26 to 30 | . 2641756 | . 4408933 | 0 | 1 |  |
| Age 31 to 35 | . 2123202 | . 4089504 | 0 | 1 |  |
| Age 36 to 40 | . 0897262 | . 2857891 | 0 | 1 |  |
| Age 41 to 55 | . 0155309 | . 1236513 | 0 | 1 |  |
| Age 46 or older | . 0007413 | . 0272174 | 0 | 1 |  |
| Less than High School | . 2179839 | . 4128765 | 0 | 1 |  |
| Less than High School, Blacks | . 2442223 | . 4296255 | 0 | 1 | 616074 |
| Less than High School, Whites | . 2129684 | . 4094056 | 0 | 1 | 3222929 |
| High School | . 2945684 | . 4558486 | 0 | 1 |  |
| High School, blacks | . 3706535 | . 4829802 | 0 | 1 | 616074 |
| High School, whites | . 2800245 | . 4490109 | 0 | 1 | 3222929 |
| More than High School | . 4744328 | . 499346 | 0 | 1 |  |
| More than High School, blacks | . 368222 | . 4823224 | 0 | 1 | 616074 |
| More than High School, whites | . 4947354 | . 4999724 | 0 | 1 | 3222929 |
| Prenatal Care from $1^{\text {st }}$ trimester | . 7987876 | . 4009065 | 0 | 1 |  |
| Prenatal Care from ${ }^{\text {st }}$ tri., blacks | . 7014498 | . 4576225 | 0 | 1 | 616074 |
| Prenatal Care from $1^{\text {st }}$ tri., whites | . 8173941 | . 3863432 | 0 | 1 | 3222929 |
| Prenatal Care from ${ }^{\text {nd }}$ trimester | . 134661 | . 3413612 | 0 | 1 |  |
| Prenatal Care from $2^{\text {nd }}$ tri., blacks | . 1924038 | . 3941889 | 0 | 1 | 616074 |
| Prenatal Care from $2^{\text {nd }}$ tri., whites | . 1236233 | . 3291513 | 0 | 1 | 3222929 |
| Prenatal Care from $3^{\text {rd }}$ trimester | . 0280992 | . 1652564 | 0 | 1 |  |
| Prenatal Care from $3^{\text {rd }}$ tri., blacks | . 043959 | . 2050041 | 0 | 1 | 616074 |
| Prenatal Care from $3^{\text {rd }}$ tri., whites | . 0250676 | . 1563304 | 0 | 1 | 3222929 |
| No prenatal care | . 0106661 | . 1027243 | 0 | 1 |  |
| No prenatal care, blacks | . 0202508 | . 1408572 | 0 | 1 | 616074 |
| No prenatal care, whites | . 0088339 | . 0935727 | 0 | 1 | 3222929 |
| Smoked | . 1087593 | . 3113371 | 0 | 1 |  |
| Smoked, blacks | . 0805394 | . 2721266 | 0 | 1 | 616074 |
| Smoked, whites | . 1087593 | . 3113371 | 0 | 1 | 3222929 |
| Physicians per 100k | 293.485 | 161.3936 | 90 | 1236 |  |
| County poverty level | . 1308798 | . 0419423 | . 041 | . 305 |  |
| Education level unknown | . 0130148 | . 1133378 | 0 | 1 |  |
| Prenatal care unknown | . 0277861 | . 1643595 | 0 | 1 |  |
| Smoked unknown | . 1276196 | . 3336658 | 0 | 1 |  |
| Neonatal per 1000 | 4.552484 | 67.31835 | 0 | 1000 |  |
| Neonatal per 1000, blacks | 8.867441 | 93.74873 | 0 | 1000 | 616074 |
| Neonatal per 1000, whites | 3.727665 | 60.94072 | 0 | 1000 | 3222929 |
| Postneonatal per 1000 | 2.292522 | 47.82538 | 0 | 1000 |  |
| Postneonatal per 1000, blacks | 4.32091 | 65.59151 | 0 | 1000 | 616074 |
| Postneonatal per 1000, whites | 1.904789 | 43.60231 | 0 | 1000 | 3222929 |

Unless otherwise noted number of observations: 3839003

## Proposed Models

### 1.1 Crude Model

Primary focus of this paper was on the difference in IMR between black and white mothers. Following crude model estimates the difference between IMR of black and white mothers.

$$
\begin{equation*}
\text { Outcome per } 1000=a+b * \text { Mother is Black }+u \tag{Equation1.1}
\end{equation*}
$$

Where Outcome is dataset dummy variable Outcome_1000 defined as 0 if infant lived to his/her first birthday and 1000 otherwise (multiplying by 1000 allows comparing my results with IMR usually stated as number of deaths per 1000 births). Mother is Black, variable mblack in dataset equal to 1 if mother's race black and 0 otherwise. Observations, where mother neither black nor white, are excluded in all regressions in this study. The term $u$ contains unobserved factors affecting outcome and errors in measuring.

| Crude model: outcome_1000 | 5.6325 | + | 7.556 mblack | (Estimate 1.1) |
| :---: | :---: | :---: | :---: | :---: |
|  | (.046) |  | (.1146) |  |
|  | 122.71 |  | 65.94 |  |

Regression shows, that possibility for black mother that her child didn't live to the first birthday increases by 7.55 per 1000, compare to the white counterpart. Result is not trivial and statistically significant with standard error . 1146 and t-statistics 65.94 with $95 \%$ confidence interval.

I assume that independent variables are exogenous due to very nature of our dependent variable outcome. I also assume that the dependent variable has a conditional expectation linear in functions of the independent variables; which allows partial effects to be correctly estimated by ordinary least squares. Latter assumption can be verified by estimating average partial effect by probit model and comparing results with ordinary least square (OLS) estimates.

I also use probit model to predict the outcome. Average partial effect estimated by probit model is 7.559 for change of binary variable mblack from 0 (non-black) to 1 (black) which is the same as ordinary least square (OLS) given in Estimate 1.1. This means we can use OLS estimates crude model.

### 1.2 Crude Model with control for the states

Crude model showed us that a difference in IMR between blacks and whites exists, is statistically significant, and is large in magnitude. We need to add more independent variables to the model and see if controlling for other factors will reduce or eliminates this difference. From policy maker perspective, a straightforward way is to control for mother's state of residence. If including state of residence dummy variable to crude model will significantly reduce the estimated difference, then states with best policies have lowest estimates. Entire problem will narrow down to adopting most effective policies in states with high estimates. Additionally, blacks might live disproportionally is states with otherwise high IMR.

Outcome per $1000=a+b *$ Mother is Black $+K_{i}$ dummy state variable $+u$
(Equation 1.2)

Where dummy state variable is separate dummy variable for each state. It equal to 1 if mother resides in that state, 0 otherwise.


As we can see coefficient on mblack is lower than in our crude model (Estimate 1.1). Controlling for the state reduced our difference by 0.2678 per $1000 .{ }^{21}$ It also shows that difference in our model largely unrelated to the state of residence.

[^4]
### 2.1 Basic Model with state and age controls

Based on the literature we know that IMR is highly affected by the age of the mother. Studies showed that having birth in teens ${ }^{22}$ or after 40 greatly increase IMR. Even though in the free society we can't possibly dictate woman when to have a baby ${ }^{23}$, at least we can control for the effect in our model. In addition, if blacks tend to have births at younger or older ages than whites then failing to control for age could lead to spurious race differentials.

Outcome per $1000=a+b^{*}$ Mother is Black $+K_{i}$ dummy state variable $+L_{i}$ dummy age group variable $+u$
(Equation 2.1)
Where $L_{i}$ dummy age group variable is separate dummy variable for age bracket. It equal to 1 if mother is in the bracket, 0 otherwise. ${ }^{24}$

Outcome_1000 $=3.8886+6.97$ mblack $+\mathrm{K}_{\mathrm{i}}$ dummy state var. $+\mathrm{L}_{\mathrm{i}}$ dummy age group var.
(Estimate 2.1)
27.94

Estimate at mblack was reduced by another 0.3182 per 1000 observations. This leads me to conclusion that more white mothers give a birth in preferable age range, compared to blacks. ${ }^{25}$

To prove existence of a preferable maternal age, I used my entire dataset to estimate simple model for U.S. population in 2004. I created mean IMR for every age in the dataset as depend variable.

Outcome per 1000 mean $=a+b^{*}$ mother's age $+c$ mother's age ^ $2+u$
(Equation 2.2)

[^5]I assumed mother's age has diminishing effect on mean IMR model it as quadratic function.

| IMR 1000_agemean $=$ | 37.6 | - | $2.072 \mathrm{MAGER}+$ |
| ---: | :--- | :--- | :--- |
|  | $(3.43)$ | $(0.24)$ | 0.0329 mager_sq |
| 10.96 | -8.62 | $(.0038)$ |  |
|  |  | 8.63 |  |

Figure 1. IMR by Age. Actual values and regression line.


Figure 1, shows that data follows quadratic model fairly close. Actual values have high variance after age of 45 and were slightly under predicted by the model in early teens and over predicted in the late teens. Overall model shows a good fit. ${ }^{26}$ Based on the data, IMR below 10 in 2004 was from age 16 to 45 with the lowest points between ages 27 to 35 .

[^6]
### 3.1 Education

Since SDC datasets does not contain information about income and health coverage of the mother, I will use level of education attained by mother as my proxy for level of income. The idea is that higher level of education leads to a higher income level. I want to test whether higher level of income leads to lower IMR.

Model with controlled for the state and age (Equation 2.1) will be used as basic model for my paper. I'll add more variables to my basic model to evaluate whether controling for them reduces difference in IMR between balcks and whites.

Let's add education level to our basic model and estimate it for blacks, whites and both group combined. ${ }^{27}$

Outcome per $1000=a+b^{*}$ Mother is Black $+K_{i}$ dummy state variable $+L_{i}$ dummy age group variable + $E_{i}$ dummy education level variable $+u$ (Equation 3.1)

Table 3.1 Basic model with control for education level with less than High School as a base group

| Section 3.1 | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | 6.726 | .12 |  |  |  |  |
| high school | -.6 | .124 | -.246 | .405 | -.66 | .124 |
| more than HS | -2.5 | .123 | -2.62 | .43 | -2.48 | .123 |
| educ unknown | 17.17 | .382 | 28.53 | 1.17 | 14.25 | .389 |
| cons | 4.77 | .163 | 11.45 | .77 | 4.73 | .156 |

Coefficient on black is down to 6.73 comparing to 6.97 (Estimate 2.1). Thus controlling for education reduces our difference in IMR between blacks and whites. Thus higher education level improves IM rate. Estimates show that with each level of education IMR gets smaller for combined group and whites. Attained level of more than high school reduces IMR by 2.5 for 1000.

[^7]For blacks, there is small, statistically insignificant difference in IMR between mothers with high school level of education and less than high school, which is about $1 / 3$ of the effect for whites. However there is a drop in IMR for black mothers with more than high school. Dzietham-Picciotto (1998) in their paper noted higher education to appear ineffective in reducing $I M R$ for blacks ${ }^{28}$. One of the possible explanation is (if we assume that higher education leads to higher income leads to lower IMR), that the difference in income between black mother with high school diploma and black mother with less than high school is insignificant to lower IMR. For white mothers the difference in income between high school and less than high school is significant enough to have an effect on the IMR.

### 4.1 Trimester prenatal care started

There is consensus among medical professionals that woman should start prenatal care as soon as she learns that she is pregnant. However, there are wide gap between black and white mothers regarding who has access to and who utilizes prenatal care. According to the Office of Minority Health, African American mothers were 2.5 times as likely as non-Hispanic white mothers to begin prenatal care in the 3rd trimester, or not receive prenatal care at all. One of the reasons for starting late is that black mothers do not have health coverage. In 2007, 19.5 percent of African-Americans in comparison to 10.4 percent of non-Hispanic whites were uninsured ${ }^{29}$

Outcome per $1000=a+b^{*}$ Mother is Black $+K_{i}$ dummy state variable $+L_{i}$ dummy age group variable + $C_{i}$ dummy prenatal care $+u$ (Equation 4.1)

Adding control for the trimester prenatal care begun to our basic model, gives us opportunity to test the hypothesis that not having prenatal care or starting it late increases IMR.

[^8]$\qquad$

Table 4.1 Basic model with control for prenatal care with $1^{\text {st }}$ trimester as a base group

| Section 4.1 | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | 6.51 | .12 |  |  |  |  |
| $2^{\text {nd }}$ trimester | -.172 | .126 | -1.76 | .38 | .287 | .13 |
| $3^{\text {rd }}$ trimester | -1.15 | .257 | -4.14 | .72 | -.145 | .27 |
| no precare | 27.62 | .411 | 39.64 | 1.04 | 22.3 | .445 |
| precare <br> unknown | 10.67 | .26 | 16.07 | .75 | 9.06 | .27 |
| cons | 3.66 | .14 | 10.45 | .69 | 3.55 | 0.13 |

Controlling for prenatal care reduces our coefficient on black to 6.51. Our data confirms the Office of Minority Health report and suggests that blacks either do not have access or underutilize prenatal care. Data also showed that starting prenatal care in the third trimester do not have negative impact on mortality rate. In fact it is associated with a reduction in the mortality rate for both black and whites. One of possible explanation to the paradox, in my opinion, lies in pregnant woman's self-assessment. When woman feels that pregnancy goes as it should without any complications, she might postpone her prenatal visits till the last trimester. ${ }^{30}$

There is a significant increased risk of infant mortality for women who did not use prenatal care at all (no precare category). Penalty is very high both groups and estimated coefficients among respective groups are 39 for blacks and 22 for whites. Possible explanation is lack of health education, incorrect self-assessment or imprudence. ${ }^{31}$ For factors associated with trimester prenatal care started and discussion please see Appendix B.

All in all, the model showed that not utilizing prenatal care leads to fivefold increase in IMR over 1 trimester. Perhaps, improving prenatal care attendance can significantly reduce IMR for both blacks and whites. In this area, appropriate public policy can have greatest effect on IMR in U.S.

[^9]
### 4.2 Health Care Accessibility

It's reasonable to ask whether deficiency in health care services, such us shortage of doctors and hospitals in the area, lead to higher IMR. I matched my mother's residence county codes in main datasets with U.S Census Bureau, County and City Data Book numbers on physicians and hospital beds per 100000 population, per county. ${ }^{32}$

The next two estimates have a constraint on results interpretation. My main individual level dataset taken from CDC data does not show mother residency for low populated counties. If a woman resides in the small county, instead of Federal Information Processing Standard ${ }^{33}$ (FIPS) county code CDC listed code 999. It was done to protect the privacy. To alleviate this I used county data for counties listed in CDC data (usually metropolitan or highly populated counties) and for small areas (999 codes) I used averages for the rest of counties of that particular state. For example, Providence County, Rhode Island is listed in CDC data, and I used figures given in County and City Data Book. Other four counties: Bristol, Kent, Newport and Washington are relatively small and listed under code 999 in CDC data. In my estimates I used average number of physicians/hospital beds for those four counties, when mother's residence was in one of four counties. I applied same principal for all states.

Outcome per $1000=a+b^{*}$ Mother is Black $+K_{i}$ dummy state variable $+L_{i}$ dummy age group variable $+F_{i}$ number of physicians per 1000 per county $+u$ (Equation 4.2.1)

Table 4.2.1 Basic model with control for number of physicians per 1000 capita per county

| Section 4.2.1 | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficie <br> nt | SE | Coefficient | SE | Coefficient | SE |
| mblack | 7.045 | .12 |  |  |  |  |
| Physicians per 1000 | -.137 | .029 | -.0977 | .091 | -.14 | .03 |
| cons | 4.26 | .16 | 11.13 | .74 | 4.2 | .157 |

[^10]Outcome per $1000=a+b^{*}$ Mother is Black $+K_{i}$ dummy state variable $+L_{i}$ dummy age group variable + $H_{i}$ number of hospital beds per 1000 per county $+u$ (Equation 4.2.2)

Table 4.2.2 Basic model with control for number of hospital beds per 1000 capita per county

| Section 4.2.2 | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | 6.957 | .122 |  |  |  |  |
| Beds per 1000 | .037 | .05 | .0315 | 0.135 | .031 | .054 |
| cons | 3.82 | .17 | 10.78 | .746 | 3.74 | .17 |

Coefficients in both models estimated relatively small effect, possibly due to above mentioned constrained. Models showed that we could use better controls for the level and proximity of health services to properly estimate the effect more accurately.

Number of physicians per 1000 population is small but statistically positive effect for whites. At the same time coefficient is statistically insignificant for blacks. There is more than one explanation why only whites can benefit from additional doctors in the area. For example: underutilization of health services by blacks; blacks live in counties with disproportionally high rate of physicians per capita.

### 5.1 Smoking during pregnancy

Smoking during pregnancy was long known for its negative impact on fetal development. According to Robert Welch, the chairman of the Department of Obstetrics and Gynecology at Providence Hospital in Southfield, Michigan, "Smoking cigarettes is probably the No. 1 cause of adverse outcomes for babies". Cigarette smoke contains more than 4,000 chemicals, including cyanide, lead, and at least 60 cancercausing compounds. When a woman smokes during the pregnancy, toxins get into mother's bloodstream and transfer to the baby through the placenta. There is large number of medical papers were written about harmful effect of smoking on unborn child. More than 100 years ago John Williams Ballantyne, considered one of the founding fathers of modern prenatal care ${ }^{34}$, noted that miscarriages were more common in female tobacco factory workers.

34 John Williams Ballantyne biography can be found on http://humupd.oxfordjournals.org/content/5/4/386.full.pdf

Cramer (1995) in his study noted that minorities smoke less during the pregnancy, depressing differences in birthweight. Thus controlling for the tobacco use during the pregnancy seems to be viable option. Perhaps, smoking habits can be adjusted (at least for the time of pregnancy), through the proactive public policy and education of the smokers. ${ }^{35}$

Outcome per $1000=a+b^{*}$ Mother is Black $+K_{i}$ dummy state variable $+L_{i}$ dummy age group variable $+S_{i}$ smoked during pregnancy $+u$ (Equation 5.1)

Table 5.1 Basic model with control for smoking

| Section 5.1 | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | 7.12 | .12 |  |  |  |  |
| smoked | 4.47 | .15 | 7.23 | .557 | 3.95 | .145 |
| smoked <br> unknown | 14.48 | .62 | 32.17 | 1.98 | 10.54 | .62 |
| cons | 3.63 | .32 | 6.82 | 1.81 | 3.84 | .297 |

As we can see smoking raises IMR by more than 7 for blacks and almost 4 for whites. At the same time controlling for smoking increased difference between blacks and whites. Assuming negative influence of smoking and that mother data implies that white mothers smoke more often during the pregnancy. ${ }^{36}$

### 6.1 Poverty Level

From individual level data provided by CDC we can't tell if mother of the child lives in poverty. I already tried to address this by linking income to mother's level of education attained. In this section I want to check if combining individual records with poverty level records by county, taken from U.S. Census Bureau, will reduce the gap between black and whites. In my dataset county poverty level is percentage of population living in poverty in that county divided by hundred. To merge records I used the same method as in the health care accessibility section.

[^11]Outcome per $1000=a+b^{*}$ Mother is Black $+K_{i}$ dummy state variable $+L_{i}$ dummy age group variable + $V_{i}$ county poverty level $+u$ (Equation 6.1)

Table 6.1 Basic model with control for county poverty level

| Section 6.1 | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | 6.94 | .12 |  |  |  |  |
| county poverty | 2.104 | 1.27 | 2.695 | 4.127 | .63 | 1.3 |
| cons | 3.607 | .22 | 10.468 | 10.47 | 3.71 | .22 |

Coefficient on county poverty level is insignificant with $5 \%$ confidence interval, while significant with 10\% confidence interval.

It gives us weak evidence that higher poverty level in the county leads to higher IMR. But not knowing income on individual level blurs results. Mario Sims ${ }^{37}$ et al (2007) tried to address this issue by comparing $I M R$ in metropolitan areas, but results were statistically insignificant. Aber et al ${ }^{38}$ (1997) precisely pointed out the issue that "few useful socioeconomic covariates appear on birth or death certificates". Lack of such information seriously hinders the ability of the researcher to produce reliable estimates. One of the ways to improve the results quality and prove straightforward relationship between IMR and poverty is to conduct an experiment ${ }^{39}$.

### 6.2 Marital Status

Here I attempted to link IMR with social economic status, without depending on income level. Even though income level is lacking in individual level data, marital status is still present. Idea behind it is simple. Use marital status as a double proxy. First it shows lower bound of family income, second it reveals amount of time and support can be given to pregnant woman and later to a newborn.

[^12]The lower bound in family income is based on the assumption that even when a woman doesn't work, her husband has a job. ${ }^{40}$ After all if you're single you can't make more than single income. A married couple can have the single income of the husband plus the income of the wife multiplied by probability of her having a job. The amount of time and support given to woman and newborn follows from the first assumption. Even though, both assumptions are weakened by the possibility that pregnant woman might still live with her parents, I believe it worth to try to accommodate marital status into the model.

Outcome per $1000=a+b^{*}$ Mother is Black $+K_{i}$ dummy state variable $+L_{i}$ dummy age group variable + $P_{i}$ married $+u$ (Equation 6.2)

Table 6.2 Basic model with control for marital status

| Section 6.2 | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | 6.21 | .12 |  |  |  |  |
| Married | -2.45 | .1 | -3.65 | .35 | -2.14 | .1 |
| cons | 5.72 | .15 | 12.53 | .71 | 5.4 | .15 |

Result showed that controlling for marital status significantly reduced the difference between blacks and whites. No other socio-economic variable evaluated in this paper was able to shrink the gap to such degree. Mentioned earlier possibility that woman could still live with her parents, which will lessen income effect and make results insignificant, proven to be wrong. In fact results showed that true influence of marital status on IMR is even higher, because at least some single women are sheltered by her parents. Thus we can anticipate the real impact of being married on IMR even is higher.

Perhaps, having less income due to being single and lack of support ${ }^{41}$ puts woman in considerable disadvantage and leads to higher IMR. The impact is very large and significant for both groups. The model implies that black mothers more often to have children while being single compare to white mothers, and hence suffer from adverse effect more often. ${ }^{42}$

[^13]
### 7.1 Combined Model

Based on discussions in previous sections I believe following variables belong to the model that explains gap in IMR between blacks and whites. Education attained, trimester prenatal care begun, county poverty level and physicians per capita, smoking habits and marital status. Most of them have positive impact on reducing gap between black and whites.

Outcome per $1000=a+b^{*}$ Mother is Black $+K_{i}$ State $+L_{i}$ Age Group $+E_{i}$ Education Level $+P_{i}$ Married $+C_{i}$ Prenatal Care $+V_{i}$ County Poverty Level $+F_{i}$ physicians per $1000+S_{i}$ smoked $+u$ (Equation 7.1)

Table 7.1 Combined factors model

| Section 7.1 | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | 6.22 | .13 |  |  |  |  |
| high school | -.14 | .12 | .83 | .41 | -.3 | 0.125 |
| more than HS | -1.39 | .13 | -.44 | .44 | -1.52 | 0.13 |
| $2^{\text {nd }}$ trimester | -.78 | .13 | -2.34 | .38 | -.31 | 0.13 |
| $3^{\text {rd }}$ trimester | -1.98 | .26 | -4.99 | .72 | -.95 | 0.27 |
| no precare | 25.93 | .41 | 37.14 | 1.05 | 20.83 | 0.45 |
| married | -1.33 | .11 | -2.37 | .36 | -1.08 | .11 |
| smoked | 3.54 | .15 | 5.41 | .57 | 3.18 | .15 |
| county poverty | -.73 | 1.28 | -0.94 | 4.14 | -1.54 | 1.31 |
| Docs per 1000 | -.15 | .03 | -.176 | .09 | -.13 | 0.03 |
| smoked <br> unknown | 8.88 | .62 | 22.45 | 1.99 | 6.01 | .63 |
| precare <br> unknown | 8.96 | .26 | 12.9 | 0.76 | 7.7 | .27 |
| educ unknown | 15.39 | .38 | 24.37 | 1.19 | 13.02 | .39 |
| cons | -3.27 | .67 | -10.86 | 2.23 | -.48 | .68 |

Our coefficient on mblack fell to 6.22, compare to 7.55 in crude model (Estimate 1.1) and 6.97 in basic model (Estimate 2.1). (Full results available in Appendix A tables 7.1a-c). Thus, selected socioeconomics factors able to explain only 0.75 out 6.97 deaths difference between blacks and whites or about $10.76 \%$. All other coefficients in column one (black and white) behave in expected way described in previous sections.

To better understand the differences between black and whites let's concentrate on columns two and three: black only and white only.

The education level, (as we discussed in section 3.1 it's our proxy for income) gradually decreases IMR for whites. For blacks, we start seeing decrease only at education level more than high school, and decrease three times lower compare to whites. One the possible explanations is that blacks can't use attained level of education to generate income at the same rate as whites do (possibly due to acquiring less human capital in school for any reason including quality of teaching and incentive to learn).

For whites we can see gradual improvement towards $3^{\text {rd }}$ trimester, but it doesn't mean that it's better to start prenatal care in the last three months. As I proposed in section 4.1 paradoxical improvements in IMR could be due to self-assessment effect. For blacks self-assessment effect more than 5 times higher than for whites. When we look at coefficient on no prenatal care variable, it's more than 37.1 for blacks and only 20.8 for whites (these two numbers are very high in magnitude and largest increase in IMR for both groups). Section 4.1 explains it as a combination of self-assessments and imprudence.

Being married is very important factor in lowering IMR for both groups, but for blacks the positive effect is twice as higher. As we discussed in section 6.2 being married helps lower IMR in two ways: additional income and support.

If I exclude smoking related variables from the model, model will explain 1.68 deaths difference or $22 \%$. It demonstrates that blacks smoke less during the pregnancy which helps them narrow the gap.

The county poverty level was borderline significant with $10 \%$ confidence interval (Table 6.1), but in combined model changed sign and lost any significance due to high positive correlation with variables black, mother education more than high school, married (negative correlation). ${ }^{43}$

For both groups, coefficients on unknown (variables with missing data) are associated with adverse levels of the IMR. I can't find a reasonable explanation, but I decided to keep them in the model because removing records with missing data significantly reduces sample size and disproportionally removes
${ }^{43}$ I could drop county poverty level from final combined model but effect of not omitting it is negligible.
blacks from the sample. ${ }^{44}$ Another reason is that correlation between different unknowns is not as high as we might think. It's common to have missing data only for one factor, for example, trimester prenatal care started, but have data for two other factors: mother education and smoking during the pregnancy. ${ }^{45}$

### 7.2 Combined Model for Neonatal Mortality

By convention, infant mortality analysis often subdivided into two periods-the neonatal and postneonatal. Neonatal period or first 28 days of life ${ }^{46}$ and postneonatal period covers the remaining eleven months of the first year of life.

Based on data out of 26,278 deaths, neonatal mortality accounted for 17,477 occurrences or $66.51 \%$ and postneonatal accounted for 8,801 occurrences or $33.49 \%$ occurrences. Among blacks neonatal mortality was $67.24 \%$ and among whites $66.18 \%{ }^{47}$

The skewness in infant mortality towards first month of life suggests that factors influencing IMR in neonatal and postneonatal period are different or at least magnitude of factors changes over periods. Based on this assumption we need to check how factors of our model are going to behave in each period.

Neonatal Mortality per 1000 $=a+b^{*}$ Mother is Black $+K_{i}$ State $+L_{i}$ Age Group $+E_{i}$ Education Level $+P_{i}$ Married $+C_{i}$ Prenatal Care $+V_{i}$ County Poverty Level $+F_{i}$ physicians $+S_{i}$ smoked $+u$ (Equation 7.2)

[^14]Table 7.2 Combined factors Model for Neonatal Period

| Section 7.2 Blacks and Whites | Blacks Only |  | Whites Only |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | 4.34 | .1 |  |  |  |  |
| high school | .3 | .1 | 1.87 | .34 | .008 | 0.1 |
| more than HS | -.47 | .1 | 1.27 | .36 | -.74 | 0.1 |
| $2^{\text {nd }}$ trimester | -1.1 | .1 | -2.72 | .31 | -.64 | .11 |
| $3^{\text {rd }}$ trimester | -2.36 | .21 | -5.44 | .59 | -1.34 | .22 |
| no precare | 21.88 | .34 | 31.86 | .86 | 17.39 | .36 |
| married | -.82 | .09 | -1.47 | .29 | -.64 | .086 |
| smoked | 1.10 | .12 | 1.51 | .47 | 1.00 | .12 |
| county poverty | -1.2 | 1.04 | -3.74 | 3.43 | -1.36 | 1.07 |
| Docs per 1000 | -.11 | .024 | -.12 | .075 | -.1 | .025 |
| smoked <br> unknown | 7.24 | .51 | 18.05 | 1.64 | 4.93 | .51 |
| precare <br> unknown | 7.81 | .22 | 10.9 | .62 | 6.78 | .22 |
| educ unknown | 15.52 | .31 | 25.99 | .97 | 12.82 | .32 |
| cons | -3.73 | .55 | -11.26 | 1.83 | -1.39 | .55 |

Coefficients estimated for neonatal period amplify behavior of our combined model (Table 7.1), full results in Appendix A. Tables 7.2a-c. In neonatal period only whites get reduction in IMR with increased level of mother's education, and only when mother's education higher than high school. Both groups have seemingly lower IMR when they postpone beginning of prenatal care. I gave possible explanation to this paradox in Section 4.1, and we will continue to discuss in Section 7.3

### 7.3 Combined Model for Postneonatal Mortality

Following model conditioned on infant survival in neonatal period.
Postneonatal Mortality per 1000 $=a+b^{*}$ Mother is Black $+K_{i}$ State $+L_{i}$ Age Group $+E_{i}$ Education Level + $P_{i}$ Married $+C_{i}$ Prenatal Care $+V_{i}$ County Poverty Level $+F_{i}$ physicians $+S_{i}$ smoked $+u$ (Equation 7.3)

Table 7.3 Combined factors Model for Postneonatal Period

| Section 7.3 | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | 1.91 | .07 |  |  |  |  |
| high school | -.44 | .07 | -1.04 | .24 | -.31 | .07 |
| more than HS | -.93 | .07 | -1.71 | .26 | -.78 | .07 |
| $2^{\text {nd }}$ trimester | .32 | .07 | .376 | .22 | .33 | .077 |
| $3^{\text {rd }}$ trimester | .37 | .15 | .42 | .42 | .38 | .158 |
| no precare | 4.25 | .24 | 5.67 | .62 | 3.57 | .26 |
| married | -.52 | .06 | -.92 | .21 | -.44 | .06 |
| smoked | 2.45 | .088 | 3.95 | .33 | 2.2 | .087 |
| county poverty | .47 | .74 | 2.8 | 2.4 | -.185 | .768 |
| Docs per 1000 | -.0386 | .017 | -.057 | .053 | -.035 | .0176 |
| smoked <br> unknown | 1.7 | .36 | 4.65 | 1.17 | 1.1 | .37 |
| precare <br> unknown | 1.18 | .15 | 2.1 | .44 | .95 | .16 |
| educ unknown | -.057 | .23 | -1.49 | .7 | .25 | .23 |
| cons | .41 | .39 | .18 | 1.31 | .89 | .4 |

As we can see applying model to two periods separately yields surprising results. Coefficients estimated for postneonatal period different from estimates in combined model (Table 7.1). In postneonatal period the elements that made white and black groups behave differently disappeared. Estimates follow expected theoretical path. Gradual decrease in IMR with level of education became more prominent and significant for whites and became new pattern for blacks. Gradual increase in IMR with trimester prenatal care started reversed its course from neonatal period and became new norm for both blacks and whites. In postneonatal period everything progresses according to common sense.

Striking difference between neonatal and posneonatal periods can be explained by the different nature of mortality. Education and prenatal care variables measure different factors in different periods.

The posneonatal mortality depends on nutrition, environmental risk factors and health care coverage of infant. All three factors correlate with income. And income correlates mother's level of education and health insurance coverage. Coefficients on education are a proxy for income and trimester prenatal care began proxy for health insurance coverage.

Neonatal mortality depends on maternal health, genetic disorders, level of prenatal care, nutrition of the mother, proficiency in medical treatment and sophistication of equipment. Simply put neonatal mortality depends more on factors outside of our control than postneonatal mortality.

Education level in our model is proxy for income ${ }^{48}$, and as mentioned earlier blacks unable to turn education to income at the same rate as whites do. But more research needed to explain why black mothers with less than high school education level have lower IMR in neonatal period than black mothers with high school or more education level attained. ${ }^{49}$

In general there is a good chance to have a healthy baby without any medical intervention during the pregnancy. Most of the doctor's visits in the course of pregnancy aimed at monitoring development of the fetus and mother's health. Mothers who do not start prenatal care from the first trimester either face high insurance deductible and trying save on doctor's visits (we could think of them as risk-takers who trusts their self-assessment), imprudent mothers and mothers with no health coverage. Thus, trimester prenatal care began measures outcomes for such mothers. I assume risk-taker has an insurance or income to pay for visits and she will start seeing the doctor in the 3rd trimester (to save on copay and deductible) or sooner if she feels that something is wrong with the pregnancy. Thus negative estimates on $2^{\text {nd }}$ and $3^{\text {rd }}$ trimesters are from risk-takers and no-insurance mothers with healthy pregnancy. Mortality is heavily concentrated on no prenatal care category. In my opinion, those are mothers with no health insurance coverage and imprudent mothers who unfortunate enough to have unhealthy pregnancy. ${ }^{50}$

[^15]Dividing on neonatal and postneonatal periods highlighted difference in causes of infant mortality between periods. Postneonatal mortality more receptive to positive changes in socioeconomics factors, at the same time more research is needed to explain dynamics of prenatal moprtality.

## Conclusion

Society achieved exceptional decline in infant mortality rate in the last century. Most of it was due to staggering achievements in modern medicine especially in areas of maternal, prenatal and infant care. ${ }^{51}$ Efforts by both public and private sector to bring improvements in care to every single mother and infant in U.S. lead to astonishing more than $90 \%$ decline in IMR. ${ }^{52}$ From other side disparity in IMR between different social and racial groups stayed significant.

This study shows that socioeconomic factors still play significant role in IMR. Targeted public programs with focus on economically and socially disadvantaged could further reduce the IMR disparity and consequently lower nation's IMR. ${ }^{53}$

Mother's level education associated with lower IMR for blacks and whites in postneonatal period. In neonatal period it has a positive effect for white mothers with education level higher than high school. Black mothers with an education at the high school level or higher have IMR rate well above black mothers with education levels less than high school.

Early prenatal care has positive impact on postneonatal mortality. For neonatal mortality results are reversed, and more research is needed to explain it. Study showed that not having prenatal care at all increases the risk of infant mortality in both periods. As result a lower rate of infant mortality can be achieved by encouraging use of prenatal care. This is especially important for black mothers who lag behind white mothers in prenatal care utilization. Encouragement can be done in form of education in black communities and schools. We can also use subsidies or tax breaks to help mothers of both races

[^16]with prenatal care payments and deductibles. Another possible way to achieve this goal is through expansion of Medicaid program since it targets those with low income.

Because infants of married women face significantly reduced rates of mortality, another recommendation is to promote family planning during school years with an emphasis on having kids in stable relationship. This should reduce disparity between blacks and whites.

And last but not least is pursuing programs that educate women about the harm of smoking during the pregnancy. This recommendation is more directed to white mothers who have higher rates of smoking.

While infant mortality is an important public health problem, is also an important socio-economic indicator. It reveals shortcomings of society and highlights areas and social groups which require attention. Addressing the issue and leveling IMR among different groups, through the appropriate and targeted public policies, will lead to a healthier, happier, and more efficient society.

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Appendix A. Table 7.1a Combined model estimates. (Blacks \& Whites)

| $>$ reg outcom <br> $>$ $A L A R A Z ~ C O$ <br> $>$ TN TX UT <br> $>$ poverty_lev <br> Source $\quad$Model <br> Residual | 000 mblack married age_11_16 a DC DE FL GA HI IA ID IL IN KS VT WA WI WV WY meduc_ru_hs me physicians_per_1000 smoked_unk |  |  |  | age_19_21 age_22_25 age_ A MD ME MI MN MO MS MT NC ths precare_2trimestr pre care_unknown meduc_ru_unkn$\begin{aligned} & \text { Number of obs }=3839003 \\ & F(71,3838931)=211.36 \end{aligned}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  |  |  | Prob > F R-squared | $\begin{aligned} & =0.0000 \\ & =0.0039 \end{aligned}$ |
| Total | $2.6098 \mathrm{e}+10$ | 83900267 | 8.15403 |  | Root MSE | $=82.291$ |
| outcome_1000 | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Conf | Interval] |
| mblack | 6.218724 | . 1270305 | 48.95 | 0.000 | 5.969749 | 6.4677 |
| married | -1.334116 | . 1063545 | -12.54 | 0.000 | -1.542567 | -1.125665 |
| age_11_16 | 3.084779 | . 3512087 | 8.78 | 0.000 | 2.396422 | 3.773136 |
| age_17_18 | . 7389242 | . 2227038 | 3.32 | 0.001 | . 3024327 | 1.175416 |
| age_19_21 | . 5770845 | . 1490377 | 3.87 | 0.000 | . 2849758 | . 8691932 |
| age_22_25 | . 0548451 | . 1249073 | 0.44 | 0.661 | -. 1899687 | . 2996589 |
| age_31_35 | . 290733 | . 1234014 | 2.36 | 0.018 | . 0488707 | . 5325954 |
| age_36_40 | 1.424916 | . 1634143 | 8.72 | 0.000 | 1.10463 | 1.745202 |
| age_41_45 | 3.64785 | . 3473287 | 10.50 | 0.000 | 2.967099 | 4.328602 |
| age_46_older | 6.912975 | 1.545072 | 4.47 | 0.000 | 3.88469 | 9.941261 |
| AK | 7.802943 | 1.169031 | 6.67 | 0.000 | 5.511683 | 10.0942 |
| AL | 11.03829 | . 7178161 | 15.38 | 0.000 | 9.631395 | 12.44518 |
| AR | 10.61304 | . 7597114 | 13.97 | 0.000 | 9.124037 | 12.10205 |
| AZ | 9.839328 | . 6838411 | 14.39 | 0.000 | 8.499024 | 11.17963 |
| CO | 10.14234 | . 7085458 | 14.31 | 0.000 | 8.753614 | 11.53106 |
| CT | 9.552228 | . 757 | 12.62 | 0.000 | 8.068535 | 11.03592 |
| DC | 10.9526 | 1.146209 | 9.56 | 0.000 | 8.706068 | 13.19913 |
| DE | 11.30816 | 1.011472 | 11.18 | 0.000 | 9.325706 | 13.29061 |
| FL | 9.4778 | . 6581489 | 14.40 | 0.000 | 8.187852 | 10.76775 |
| GA | 10.50221 | . 6696681 | 15.68 | 0.000 | 9.18968 | 11.81473 |
| HI | 9.573898 | 1.256699 | 7.62 | 0.000 | 7.110813 | 12.03698 |
| IA | 9.472425 | . 7620685 | 12.43 | 0.000 | 7.978797 | 10.96605 |
| ID | 9.977777 | . 8398791 | 11.88 | 0.000 | 8.331643 | 11.62391 |
| IL | 10.24699 | . 661915 | 15.48 | 0.000 | 8.949657 | 11.54432 |
| IN | 11.06226 | . 6906021 | 16.02 | 0.000 | 9.708707 | 12.41582 |
| KS | 11.32672 | . 7592653 | 14.92 | 0.000 | 9.838584 | 12.81485 |
| KY | 10.08517 | . 7235843 | 13.94 | 0.000 | 8.666967 | 11.50337 |
| LA | 12.13096 | . 7145858 | 16.98 | 0.000 | 10.73039 | 13.53152 |
| MA | 9.032684 | . 7053818 | 12.81 | 0.000 | 7.65016 | 10.41521 |
| MD | 10.87461 | . 7089611 | 15.34 | 0.000 | 9.485073 | 12.26415 |
| ME | 10.29593 | . 9463686 | 10.88 | 0.000 | 8.441084 | 12.15078 |
| MI | 10.03665 | . 6583015 | 15.25 | 0.000 | 8.746405 | 11.3269 |
| MN | 8.179987 | . 7016484 | 11.66 | 0.000 | 6.804781 | 9.555194 |
| MO | 10.15406 | . 6969012 | 14.57 | 0.000 | 8.788159 | 11.51996 |
| MS | 11.04587 | . 7523258 | 14.68 | 0.000 | 9.571338 | 12.5204 |
| MT | 8.844728 | 1.03942 | 8.51 | 0.000 | 6.807502 | 10.88195 |
| NC | 11.47067 | . 6756718 | 16.98 | 0.000 | 10.14638 | 12.79497 |
| ND | 10.37778 | 1.161529 | 8.93 | 0.000 | 8.101224 | 12.65434 |
| NE | 10.45507 | . 816954 | 12.80 | 0.000 | 8.853868 | 12.05627 |
| NH | 9.743612 | . 9412049 | 10.35 | 0.000 | 7.898884 | 11.58834 |
| NJ | 8.773547 | . 6831357 | 12.84 | 0.000 | 7.434626 | 10.11247 |
| NM | 8.912362 | . 8172211 | 10.91 | 0.000 | 7.310637 | 10.51409 |
| NV | 7.942006 | . 7734291 | 10.27 | 0.000 | 6.426112 | 9.4579 |
| NY | 9.382233 | . 6562804 | 14.30 | 0.000 | 8.095947 | 10.66852 |
| OH | 10.25095 | . 6664377 | 15.38 | 0.000 | 8.944759 | 11.55715 |
| OK | 10.9599 | . 7370196 | 14.87 | 0.000 | 9.515364 | 12.40443 |
| OR | 9.780294 | . 7427455 | 13.17 | 0.000 | 8.324539 | 11.23605 |
| PA | 9.795056 | . 6603402 | 14.83 | 0.000 | 8.500812 | 11.0893 |
| RI | 8.387042 | . 9710304 | 8.64 | 0.000 | 6.483857 | 10.29023 |
| SC | 11.49944 | . 7224496 | 15.92 | 0.000 | 10.08346 | 12.91541 |
| SD | 11.16542 | 1.059567 | 10.54 | 0.000 | 9.088707 | 13.24213 |
| TN | 10.09314 | . 6973846 | 14.47 | 0.000 | 8.726289 | 11.45999 |
| TX | 9.366991 | . 6461224 | 14.50 | 0.000 | 8.100614 | 10.63337 |
| UT | 9.537385 | . 7307411 | 13.05 | 0.000 | 8.105159 | 10.96961 |
| VA | 10.66879 | . 6865279 | 15.54 | 0.000 | 9.32322 | 12.01436 |
| VT | 8.754159 | 1.201966 | 7.28 | 0.000 | 6.398349 | 11.10997 |
| WA | 8.025954 | . 6984293 | 11.49 | 0.000 | 6.657057 | 9.394851 |
| WI | 10.11716 | . 7065254 | 14.32 | 0.000 | 8.732394 | 11.50192 |
| WV | 10.66574 | . 851477 | 12.53 | 0.000 | 8.996872 | 12.3346 |
| WY | 12.78855 | 1.206037 | 10.60 | 0.000 | 10.42476 | 15.15234 |
| meduc_ru_hs | -. 1357634 | . 1244971 | -1.09 | 0.275 | -. 3797733 | . 1082465 |
| meduc_ru_m~s | -1.389246 | . 127609 | -10.89 | 0.000 | -1.639355 | -1.139137 |
| precare_2t~r | -. 7762637 | . 127806 | -6.07 | 0.000 | -1.026759 | -. 5257685 |
| precare_3t~r | -1.983452 | .257969 | -7.69 | 0.000 | -2.489062 | -1.477842 |
| precare_none | 25.93063 | . 412803 | 62.82 | 0.000 | 25.12155 | 26.73971 |
| smoked | 3.53712 | . 1511775 | 23.40 | 0.000 | 3.240817 | 3.833422 |
| county_pov~1 | -. 7275355 | 1.276406 | -0.57 | 0.569 | -3.229246 | 1.774175 |
| physici~1000 | -. 1518118 | . 0293322 | -5.18 | 0.000 | -. 2093019 | -. 0943217 |
| smoked_unk~n | 8.88528 | . 6215357 | 14.30 | 0.000 | 7.667092 | 10.10347 |
| precare_un~n | 8.961381 | . 2633671 | 34.03 | 0.000 | 8.445191 | 9.477571 |
| meduc_ru_u~n | 15.39269 | . 3848942 | 39.99 | 0.000 | 14.63831 | 16.14707 |
| _cons | -3.274462 | .6733763 | -4.86 | 0.000 | -4.594255 | -1.954668 |

reg outcome_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county_poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mblack==1 | mwhite==1

Table 7.1b Combined model. (Blacks)
. reg outcome_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_o1der AK $>$ AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD $>$ TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county $>$ poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mblack==1

| Source | SS | $d f$ | MS |
| ---: | ---: | ---: | :---: |
| Mode1 | 39512615.2 | 70 | 564465.931 |
| Residua1 | $\mathbf{7 . 9 7 8 3 e}+09616003$ | 12951.7746 |  |
| Tota1 | $\mathbf{8 . 0 1 7 8 e}+09616073$ | $\mathbf{1 3 0 1 4 . 4 3 9 3}$ |  |


| Number of obs | $=616074$ |
| ---: | ---: | ---: |
| $\mathrm{~F}(70,616003)$ | $=43.58$ |
| Prob $>\mathrm{F}$ | $=0.0000$ |
| R-squared | $=0.0049$ |
| Adj R-squared | $=0.0048$ |
| Root MSE | $=113.81$ |


| outcome_1000 | Coef. | Std. Err. | t | P>\| | tl | [95\% Conf. |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: | Interva1]

reg outcome_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county_poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mblack==1

Table 7.1c Combined model. (Whites)
. reg outcome_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK $>$ AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD $>$ TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county $>$ poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mwhite==1

| Source | SS | $d f$ | $M S$ |
| ---: | :--- | :--- | :---: |
| Mode1 | 40932796.7 | 70 | 584754.239 |
| Residual | $1.8010 \mathrm{e}+103222858$ | 5588.15227 |  |
| Total | $1.8051 \mathrm{e}+103222928$ | 5600.7314 |  |

$$
\begin{aligned}
\text { Number of obs } & =3222929 \\
\mathrm{~F}(70,3222858) & =104.64 \\
\text { Prob }>\mathrm{F} & =0.0000 \\
\text { R-squared } & =0.0023 \\
\text { Adj R-squared } & =0.0022 \\
\text { Root MSE } & =74.754
\end{aligned}
$$

outcome_1000 Coef. Std. Err. $\quad$ t $P>|t| \quad$ [95\% Conf. Interva1]

| mb1ack | (dropped) |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| married | -1.076634 | .1063067 | -10.13 | 0.000 | -1.284991 | -.8682768 |
| age_11_16 | 3.928351 | .3818188 | 10.29 | 0.000 | 3.179999 | 4.676702 |
| age_17_18 | 1.4146 | .2315044 | 6.11 | 0.000 | .9608598 | 1.868341 |
| age_19_21 | 1.101968 | .1511973 | 7.29 | 0.000 | .805627 | 1.39831 |
| age_22_25 | .2908802 | .1240299 | 2.35 | 0.019 | .0477861 | .5339744 |
| age_31_35 | .2300145 | .1195821 | 1.92 | 0.054 | -.0043622 | .4643912 |
| age_36_40 | 1.421012 | .1585902 | 8.96 | 0.000 | 1.110181 | 1.731843 |
| age_41_45 | 3.618073 | .338657 | 10.68 | 0.000 | 2.954317 | 4.281829 |
| age_46_o1der | 7.616078 | 1.491086 | 5.11 | 0.000 | 4.693602 | 10.53855 |
| AK | 4.40092 | 1.118284 | 3.94 | 0.000 | 2.209123 | 6.592717 |
| AL | 8.301021 | .7358526 | 11.28 | 0.000 | 6.858776 | 9.743266 |


| AL | 8.301021 | .7358526 | 11.28 | 0.000 | 6.209123 | 6.592717 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| AR | 8.142522 | .7657223 | 10.63 | 0.000 | 6.641733 | 9.743266 |
| AZ | 7.209131 | .67879 | 10.62 | 0.000 | 5.878727 | 8.54331 |
| CO | 7.094123 | .7020027 | 10.11 | 0.000 | 5.718222 | 8.470023 |


| CT | 6.094123 | . 7548148 | 8.67 | 0.000 0.000 | 5.718222 5.065719 | 8.02454 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| DC | 6.759068 | 1.691794 | 4.00 | 0.000 | 3.443211 | 10.07493 |
| DE | 8.053339 | 1.051553 | 7.66 | 0.000 | 5.992333 | 10.11435 |
| FL | 6.794794 | . 663533 | 10.24 | 0.000 | 5.494293 | 8.095295 |
| GA | 7.20568 | . 6807432 | 10.59 | 0.000 | 5.871447 | 8.539912 |
| HI | 5.932779 | 1.218928 | 4.87 | 0.000 | 3.543722 | 8.321836 |
| IA | 6.82049 | . 7485665 | 9.11 | 0.000 | 5.353326 | 8.287654 |
| ID | 7.085373 | . 8108047 | 8.74 | 0.000 | 5.496225 | 8.674522 |
| IL | 7.258263 | . 6659859 | 10.90 | 0.000 | 5.952954 | 8.563572 |
| IN | 7.990395 | . 6902565 | 11.58 | 0.000 | 6.637517 | 9.343273 |
| KS | 8.479592 | . 7497832 | 11.31 | 0.000 | 7.010043 | 9.949141 |
| KY | 7.195165 | . 718849 | 10.01 | 0.000 | 5.786246 | 8.604084 |
| LA | 9.468695 | . 7466192 | 12.68 | 0.000 | 8.005348 | 10.93204 |
| MA | 6.349789 | . 704212 | 9.02 | 0.000 | 4.969559 | 7.73002 |
| MD | 7.681597 | . 732782 | 10.48 | 0.000 | 6.24537 | 9.117824 |
| ME | 7.380855 | . 9057824 | 8.15 | 0.000 | 5.605554 | 9.156157 |
| MI | 6.852632 | . 660957 | 10.37 | 0.000 | 5.557179 | 8.148084 |
| MN | 5.86769 | . 7002066 | 8.38 | 0.000 | 4.49531 | 7.24007 |
| MO | 7.317417 | . 698543 | 10.48 | 0.000 | 5.948297 | 8.686536 |
| MS | 7.626938 | . 8036449 | 9.49 | 0.000 | 6.051823 | 9.202054 |
| MT | 6.004261 | . 9839805 | 6.10 | 0.000 | 4.075694 | 7.932828 |
| NC | 7.880206 | . 6824099 | 11.55 | 0.000 | 6.542707 | 9.217706 |
| ND | 7.491124 | 1.094758 | 6.84 | 0.000 | 5.345438 | 9.636811 |
| NE | 7.48115 | . 8002188 | 9.35 | 0.000 | 5.912749 | 9.04955 |
| NH | 6.766248 | . 9025199 | 7.50 | 0.000 | 4.997341 | 8.535155 |
| NJ | 6.194459 | . 6879312 | 9.00 | 0.000 | 4.846139 | 7.54278 |
| NM | 6.22695 | . 7927502 | 7.85 | 0.000 | 4.673187 | 7.780712 |
| NV | 5.117396 | . 7637662 | 6.70 | 0.000 | 3.620441 | 6.61435 |
| NY | 6.620419 | . 6607525 | 10.02 | 0.000 | 5.325367 | 7.91547 |
| OH | 7.304039 | . 6697349 | 10.91 | 0.000 | 5.991382 | 8.616696 |
| OK | 8.078615 | . 7325239 | 11.03 | 0.000 | 6.642894 | 9.514336 |
| OR | 6.985275 | . 7291458 | 9.58 | 0.000 | 5.556175 | 8.414375 |
| PA | 7.189832 | . 666741 | 10.78 | 0.000 | 5.883044 | 8.496621 |
| RI | 6.13065 | . 9501975 | 6.45 | 0.000 | 4.268296 | 7.993003 |
| SC | 8.255044 | . 7478524 | 11.04 | 0.000 | 6.789279 | 9.720808 |
| SD | 8.342855 | 1.005187 | 8.30 | 0.000 | 6.372725 | 10.31299 |
| TN | 7.159233 | . 7039061 | 10.17 | 0.000 | 5.779602 | 8.538864 |
| TX | 6.570013 | . 6487465 | 10.13 | 0.000 | 5.298493 | 7.841534 |
| UT | 6.56849 | . 71814 | 9.15 | 0.000 | 5.16096 | 7.976019 |
| VA | 7.572517 | . 6937852 | 10.91 | 0.000 | 6.212722 | 8.932311 |
| VT | 5.995096 | 1.129136 | 5.31 | 0.000 | 3.782029 | 8.208163 |
| WA | 5.553984 | . 6938856 | 8.00 | 0.000 | 4.193993 | 6.913975 |
| WI | 6.570988 | . 7043368 | 9.33 | 0.000 | 5.190513 | 7.951463 |
| WV | 7.930076 | . 8253948 | 9.61 | 0.000 | 6.312331 | 9.547821 |
| WY | 10.00267 | 1.13254 | 8.83 | 0.000 | 7.782935 | 12.22241 |
| meduc_ru_hs | -. 2995061 | . 1251979 | -2.39 | 0.017 | -. 5448896 | -. 0541226 |
| meduc_ru_m~s | -1.518479 | . 1270313 | -11.95 | 0.000 | -1.767456 | -1.269502 |
| precare_2t~r | -. 3079444 | . 1310156 | -2.35 | 0.019 | -. 5647304 | -. 0511584 |
| precare_3t~r | -. 9547624 | . 2699312 | -3.54 | 0.000 | -1.483818 | -. 4257068 |
| precare_none | 20.83167 | . 4486119 | 46.44 | 0.000 | 19.95241 | 21.71093 |
| smoked | 3.184445 | . 1482706 | 21.48 | 0.000 | 2.89384 | 3.475051 |
| county_pov~1 | -1.53857 | 1.312427 | -1.17 | 0.241 | -4.11088 | 1.033739 |
| physici~1000 | -. 1344004 | . 0301075 | -4.46 | 0.000 | -. 1934101 | -. 0753908 |
| smoked_unk~n | 6.01263 | . 6264878 | 9.60 | 0.000 | 4.784736 | 7.240524 |
| precare_un~n | 7.706865 | . 2734615 | 28.18 | 0.000 | 7.17089 | 8.24284 |
| meduc_ru_u~n | 13.02209 | . 3917058 | 33.24 | 0.000 | 12.25436 | 13.78982 |
| _cons | -. 4787298 | . 6792548 | -0.70 | 0.481 | -1.810045 | . 8525857 |

reg outcome_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county_poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mwhite==1

## Correlation Matrix

|  | out~1000 | mblack | married | ru_7ths | me~ru_hs | u_mths | precar.. | $\mathrm{p} \sim 2 \mathrm{tri} \mathrm{\sim r}$ | $\sim r$ | precar~e | smok | uc. | car~n | n |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| outcome_1000 | 1.0000 |  |  |  |  |  |  |  |  |  |  |  |  |  |
| mblack | 0.0336 | 1.0000 |  |  |  |  |  |  |  |  |  |  |  |  |
| married | -0.0243 | -0.2917 | 1.0000 |  |  |  |  |  |  |  |  |  |  |  |
| meduc_ru_1~s | 0.0095 | 0.0278 | -0.2788 | 1.0000 |  |  |  |  |  |  |  |  |  |  |
| meduc_ru_hs | 0.0068 | 0.0730 | -0.1457 | -0.3412 | 1.0000 |  |  |  |  |  |  |  |  |  |
| meduc_ru_m s | -0.0199 | -0.0930 | 0.3688 | -0.5016 | -0.6140 | 1.0000 |  |  |  |  |  |  |  |  |
| precare_1t~r | -0.0174 | -0.1062 | 0.1960 | -0.1731 | -0.0352 | 0.1874 | 1.0000 |  |  |  |  |  |  |  |
| precare_2t~r | 0.0001 | 0.0740 | -0.1552 | 0.1413 | 0.0374 | -0.1511 | -0.7860 | 1.0000 |  |  |  |  |  |  |
| precare_3t~r | -0.0017 | 0.0420 | -0.0808 | 0.0811 | 0.0097 | -0.0778 | -0.3388 | -0.0671 | 1.0000 |  |  |  |  |  |
| precare_none | 0.0357 | 0.0408 | -0.0711 | 0.0627 | 0.0051 | -0.0624 | -0.2069 | -0.0410 | -0.0177 | 1.0000 |  |  |  |  |
| smoked | 0.0160 | -0.0283 | -0.1684 | 0.0986 | 0.0897 | -0.1605 | -0.0722 | 0.0585 | 0.0325 | 0.0367 | 1.0000 |  |  |  |
| meduc_ru_u n | 0.0256 | 0.0150 | -0.0234 | -0.0606 | -0.0742 | -0.1091 | -0.0534 | 0.0007 | 0.0080 | 0.0258 | -0.0129 | 1.0000 |  |  |
| precare_un~n | 0.0215 | 0.0376 | -0.0302 | 0.0082 | -0.0047 | -0.0263 | -0.3368 | -0.0667 | -0.0287 | -0.0176 | -0.0011 | 0.1047 | 1.0000 |  |
| smoked_unk~n | -0.0050 | -0.0921 | -0.0061 | 0.0736 | -0.0076 | -0.0699 | 0.0422 | -0.0362 | -0.0177 | -0.0113 | -0.1242 | 0.0705 | -0.0029 | 1.0000 |

Table 7.2a Combined factors Model for Neonatal Period (Blacks \& Whites)
. reg neonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK $>$ AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC S $>$ D TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county > _poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mblack==1 | mwhite==1

| Source | SS | $d f$ | $M S$ |
| ---: | :---: | :---: | :---: |
| Mode1 | 62146407 | 71 | 875301.507 |
| Residua1 | $1.7335 \mathrm{e}+103838931$ | 4515.65549 |  |
| Tota1 | $1.7397 \mathrm{e}+103839002$ | 4531.76014 |  |


| Number of obs | $=3839003$ |  |
| ---: | :--- | ---: | :--- |
| F $(71,3838931)$ | $=$ | 193.84 |
| Prob $>$ F | $=$ | 0.0000 |
| R-squared | $=0.0036$ |  |
| Adj R-squared | $=$ | 0.0036 |
| Root MSE | $=$ | 67.199 |


| neonata~1000 | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Conf. | Interval] |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| mb1ack | 4.338617 | . 1037328 | 41.82 | 0.000 | 4.135304 | 4.541929 |
| married | -. 8197165 | . 0868488 | -9.44 | 0.000 | -. 989937 | -. 6494959 |
| age_11_16 | 1.859921 | . 2867962 | 6.49 | 0.000 | 1.297811 | 2.422032 |
| age_17-18 | . 4062709 | . 1818594 | 2.23 | 0.025 | . 049833 | . 7627088 |
| age_19_21 | -. 2228516 | . 1217038 | -1.83 | 0.067 | -. 4613868 | . 0156837 |
| age_22_25 | -. 2911493 | . 101999 | -2.85 | 0.004 | -. 4910637 | -. 091235 |
| age_31_35 | . 2731495 | . 1007693 | 2.71 | 0.007 | . 0756452 | . 4706537 |
| age_36_40 | 1.367094 | . 1334437 | 10.24 | 0.000 | 1.105549 | 1.628639 |
| age_41_45 | 2.930473 | . 2836277 | 10.33 | 0.000 | 2.374573 | 3.486374 |
| age_46_older | 6.228665 | 1.261702 | 4.94 | 0.000 | 3.755774 | 8.701555 |
| AK | 5.663455 | . 9546279 | 5.93 | 0.000 | 3.792418 | 7.534492 |
| AL | 8.221674 | . 5861669 | 14.03 | 0.000 | 7.072808 | 9.37054 |
| AR | 8.276314 | . 6203785 | 13.34 | 0.000 | 7.060394 | 9.492234 |
| AZ | 8.135682 | . 558423 | 14.57 | 0.000 | 7.041193 | 9.230171 |
| CO | 8.60203 | . 5785968 | 14.87 | 0.000 | 7.468001 | 9.736059 |
| CT | 8.213173 | . 6181643 | 13.29 | 0.000 | 7.001593 | 9.424753 |
| DC | 8.956297 | . 9359917 | 9.57 | 0.000 | 7.121787 | 10.79081 |
| DE | 8.871104 | . 8259657 | 10.74 | 0.000 | 7.25224 | 10.48997 |
| FL | 7.4821 | . 5374428 | 13.92 | 0.000 | 6.428731 | 8.535468 |
| GA | 8.334634 | . 5468494 | 15.24 | 0.000 | 7.262828 | 9.406439 |
| HI | 8.512695 | 1.026217 | 8.30 | 0.000 | 6.501346 | 10.52404 |
| IA | 7.546421 | . 6223033 | 12.13 | 0.000 | 6.326728 | 8.766113 |
| ID | 8.247147 | . 6858433 | 12.02 | 0.000 | 6.902919 | 9.591376 |
| IL | 8.131311 | . 5405182 | 15.04 | 0.000 | 7.071915 | 9.190708 |
| IN | 9.047419 | . 563944 | 16.04 | 0.000 | 7.942109 | 10.15273 |
| KS | 8.833446 | . 6200142 | 14.25 | 0.000 | 7.61824 | 10.04865 |
| KY | 7.865417 | . 5908772 | 13.31 | 0.000 | 6.707319 | 9.023515 |
| LA | 8.999069 | . 583529 | 15.42 | 0.000 | 7.855373 | 10.14277 |
| MA | 7.873598 | . 5760131 | 13.67 | 0.000 | 6.744632 | 9.002563 |
| MD | 9.040489 | . 5789359 | 15.62 | 0.000 | 7.905795 | 10.17518 |
| ME | 8.7049 | . 7728023 | 11.26 | 0.000 | 7.190235 | 10.21957 |
| MI | 8.244546 | . 5375674 | 15.34 | 0.000 | 7.190933 | 9.298159 |
| MN | 6.831228 | . 5729644 | 11.92 | 0.000 | 5.708238 | 7.954218 |
| MO | 8.306798 | . 5690878 | 14.60 | 0.000 | 7.191406 | 9.42219 |
| MS | 8.367992 | . 6143474 | 13.62 | 0.000 | 7.163893 | 9.572092 |
| MT | 6.757957 | . 8487875 | 7.96 | 0.000 | 5.094363 | 8.42155 |
| NC | 9.412575 | . 5517519 | 17.06 | 0.000 | 8.33116 | 10.49399 |
| ND | 9.011086 | . 9485021 | 9.50 | 0.000 | 7.152056 | 10.87012 |
| NE | 8.82666 | . 6671227 | 13.23 | 0.000 | 7.519123 | 10.1342 |
| NH | 8.600356 | . 7685856 | 11.19 | 0.000 | 7.093956 | 10.10676 |
| NJ | 7.410279 | . 5578469 | 13.28 | 0.000 | 6.316919 | 8.503639 |
| NM | 7.065297 | . 6673407 | 10.59 | 0.000 | 5.757333 | 8.373262 |
| NV | 6.615247 | . 6315803 | 10.47 | 0.000 | 5.377372 | 7.853122 |
| NY | 7.925295 | . 535917 | 14.79 | 0.000 | 6.874917 | 8.975674 |
| OH | 8.046999 | . 5442114 | 14.79 | 0.000 | 6.980364 | 9.113634 |
| OK | 8.592394 | . 6018484 | 14.28 | 0.000 | 7.412792 | 9.771995 |
| OR | 8.023934 | . 6065242 | 13.23 | 0.000 | 6.835168 | 9.2127 |
| PA | 8.057719 | . 5392322 | 14.94 | 0.000 | 7.000843 | 9.114595 |
| RI | 7.277028 | . 7929411 | 9.18 | 0.000 | 5.722892 | 8.831165 |
| SC | 9.390156 | . 5899506 | 15.92 | 0.000 | 8.233874 | 10.54644 |
| SD | 9.360668 | . 8652398 | 10.82 | 0.000 | 7.664829 | 11.05651 |
| TN | 7.971773 | . 5694826 | 14.00 | 0.000 | 6.855608 | 9.087939 |
| TX | 7.537803 | . 527622 | 14.29 | 0.000 | 6.503683 | 8.571924 |
| UT | 7.879768 | . 5967214 | 13.21 | 0.000 | 6.710215 | 9.049321 |
| VA | 8.56255 | . 560617 | 15.27 | 0.000 | 7.463761 | 9.66134 |
| VT | 7.174178 | . 9815223 | 7.31 | 0.000 | 5.25043 | 9.097927 |
| WA | 6.286841 | . 5703356 | 11.02 | 0.000 | 5.169003 | 7.404678 |
| WI | 8.132017 | . 5769469 | 14.09 | 0.000 | 7.001221 | 9.262812 |
| WV | 8.803772 | . 695314 | 12.66 | 0.000 | 7.440981 | 10.16656 |
| WY | 10.44607 | . 984847 | 10.61 | 0.000 | 8.515807 | 12.37634 |
| meduc_ru_hs | . 3047736 | . 101664 | 3.00 | 0.003 | . 1055157 | . 5040315 |
| meduc_ru_m~s | -. 4670267 | . 1042052 | -4.48 | 0.000 | -. 6712652 | -. 2627881 |
| precare_2t~r | -1.100918 | . 1043661 | -10.55 | 0.000 | -1.305471 | -. 8963639 |
| precare_3t~r | -2.363312 | . 2106569 | -11.22 | 0.000 | -2.776192 | -1.950432 |
| precare_none | 21.88459 | . 3370939 | 64.92 | 0.000 | 21.22389 | 22.54528 |
| smoked | 1.102088 | . 1234512 | 8.93 | 0.000 | . 8601284 | 1.344048 |
| county_pov~1 | -1.205053 | 1.04231 | -1.16 | 0.248 | -3.247943 | . 8378383 |
| physici~1000 | -. 1136036 | . 0239526 | -4.74 | 0.000 | -. 1605499 | -. 0666573 |
| smoked_unk~n | 7.237216 | . 5075446 | 14.26 | 0.000 | 6.242446 | 8.231985 |
| precare_un~n | 7.815255 | . 2150649 | 36.34 | 0.000 | 7.393736 | 8.236775 |
| meduc_ru_u~n | 15.5176 | . 3143037 | 49.37 | 0.000 | 14.90158 | 16.13363 |
| _cons | -3.730403 | . 5498774 | -6.78 | 0.000 | -4.808143 | -2.652663 |

reg neonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county_poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mblack==1 \| mwhite==1

## 7.2b Combined factors Model for Neonatal Period (Blacks)

. reg neonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK $>$ AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC S $>$ D TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county $>$ _poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mblack==1

| Source | SS | $d f$ | MS |
| ---: | ---: | ---: | ---: |
| Mode1 | $\mathbf{3 0 2 0 6 5 7 3 . 2}$ | $\mathbf{7 0}$ | $\mathbf{4 3 1 5 2 2 . 4 7 4}$ |
| Residua1 | $\mathbf{5 . 3 8 4 4 e + 0 9 6 1 6 0 0 3}$ | 8740.78632 |  |
| Total | $\mathbf{5 . 4 1 4 6 e + 0 9 6 1 6 0 7 3}$ | $\mathbf{8 7 8 8 . 8 2 4}$ |  |


| Number of obs | $=616074$ |
| :--- | ---: |
| $\mathrm{~F}(70,616003)$ | $=49.37$ |
| Prob $>\mathrm{F}$ | $=0.0000$ |
| $\mathrm{R}-$ squared | $=0.0056$ |
| Adj R-squared | $=0.0055$ |
| Root MSE | $=93.492$ |


| neonata~1000 | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Con | Interval] |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| mblack | (dropped) |  |  |  |  |  |
| married | -1.46811 | . 2956921 | -4.96 | 0.000 | -2.047657 | -. 8885628 |
| age_11_16 | . 8737941 | . 7538724 | 1.16 | 0.246 | -. 6037715 | 2.35136 |
| age_17_18 | -. 4880723 | . 5347854 | -0.91 | 0.361 | -1.536235 | . 5600899 |
| age_19_21 | -1.777541 | . 3887966 | -4.57 | 0.000 | -2.53957 | -1.015512 |
| age_22_25 | -1.575376 | . 3544675 | -4.44 | 0.000 | -2. 270121 | -. 8806312 |
| age_31_35 | . 798885 | . 4047815 | 1.97 | 0.048 | . 0055262 | 1.592244 |
| age_36_40 | 1.208708 | . 529876 | 2.28 | 0.023 | . 170168 | 2.247248 |
| age_41_45 | 2.869137 | 1.087249 | 2.64 | 0.008 | . 7381634 | 5.00011 |
| age_46_older | 2.3887 | 5.201601 | 0.46 | 0.646 | -7.806271 | 12.58367 |
| AK | 19.23138 | 4.992713 | 3.85 | 0.000 | 9.445824 | 29.01694 |
| AL | 20.12399 | 1.848657 | 10.89 | 0.000 | 16.50068 | 23.74729 |
| AR | 18.52363 | 2.026274 | 9.14 | 0.000 | 14.55219 | 22.49506 |
| AZ | 16.43224 | 2.32835 | 7.06 | 0.000 | 11.86875 | 20.99573 |
| CO | 21.0536 | 2.435842 | 8.64 | 0.000 | 16.27943 | 25.82777 |
| CT | 19.97234 | 2.148207 | 9.30 | 0.000 | 15.76192 | 24.18275 |
| DC | 19.97382 | 2.185201 | 9.14 | 0.000 | 15.6909 | 24.25674 |
| DE | 21.1479 | 2.438245 | 8.67 | 0.000 | 16.36902 | 25.92678 |
| FL | 18.54249 | 1.765825 | 10.50 | 0.000 | 15.08152 | 22.00345 |
| GA | 20.95815 | 1.766089 | 11.87 | 0.000 | 17.49667 | 24.41963 |
| HI | 23.98996 | 4.270302 | 5.62 | 0.000 | 15.62031 | 32.35961 |
| IA | 14.4038 | 2.973077 | 4.84 | 0.000 | 8.576667 | 20.23094 |
| ID | 10.86563 | 9.078732 | 1.20 | 0.231 | -6.928388 | 28.65966 |
| IL | 19.97025 | 1.788474 | 11.17 | 0.000 | 16.4649 | 23.4756 |
| IN | 22.07014 | 1.95614 | 11.28 | 0.000 | 18.23617 | 25.90411 |
| KS | 18.88684 | 2.450118 | 7.71 | 0.000 | 14.08469 | 23.689 |
| KY | 19.57836 | 2.17214 | 9.01 | 0.000 | 15.32103 | 23.83568 |
| LA | 21.39076 | 1.818706 | 11.76 | 0.000 | 17.82615 | 24.95536 |
| MA | 19.13229 | 2.007031 | 9.53 | 0.000 | 15.19857 | 23.066 |
| MD | 21.34105 | 1.817174 | 11.74 | 0.000 | 17.77945 | 24.90266 |
| ME | 25.48564 | 6.464324 | 3.94 | 0.000 | 12.81578 | 38.15551 |
| MI | 20.95451 | 1.796402 | 11.66 | 0.000 | 17.43362 | 24.4754 |
| MN | 13.55462 | 2.027317 | 6.69 | 0.000 | 9.581142 | 17.52809 |
| MO | 19.49981 | 1.921914 | 10.15 | 0.000 | 15.73292 | 23.2667 |
| MS | 21.14037 | 1.854233 | 11.40 | 0.000 | 17.50613 | 24.77461 |
| MT | 12.45858 | 12.95575 | 0.96 | 0.336 | -12.93427 | 37.85144 |
| NC | 23.4581 | 1.799179 | 13.04 | 0.000 | 19.93176 | 26.98443 |
| ND | 11.89435 | 9.695714 | 1.23 | 0.220 | -7.108935 | 30.89764 |
| NE | 21.23227 | 2.91734 | 7.28 | 0.000 | 15.51437 | 26.95016 |
| NH | 29.64609 | 6.092762 | 4.87 | 0.000 | 17.70447 | 41.5877 |
| NJ | 18.44136 | 1.837677 | 10.04 | 0.000 | 14.83957 | 22.04315 |
| NM | 15.26236 | 4.503015 | 3.39 | 0.001 | 6.436591 | 24.08812 |
| NV | 18.40718 | 2.413549 | 7.63 | 0.000 | 13.6767 | 23.13765 |
| NY | 19.39136 | 1.773317 | 10.94 | 0.000 | 15.91571 | 22.867 |
| OH | 19.8513 | 1.81378 | 10.94 | 0.000 | 16.29635 | 23.40624 |
| OK | 20.946 | 2.169679 | 9.65 | 0.000 | 16.6935 | 25.1985 |
| OR | 16.1583 | 3.336928 | 4.84 | 0.000 | 9.618025 | 22.69857 |
| PA | 18.40408 | 1.763649 | 10.44 | 0.000 | 14.94739 | 21.86077 |
| RI | 12.80577 | 3.156146 | 4.06 | 0.000 | 6.61982 | 18.99171 |
| SC | 22.16032 | 1.837956 | 12.06 | 0.000 | 18.55799 | 25.76266 |
| SD | 19.45796 | 7.899717 | 2.46 | 0.014 | 3.974767 | 34.94115 |
| TN | 20.55618 | 1.859901 | 11.05 | 0.000 | 16.91084 | 24.20153 |
| TX | 18.43406 | 1.770508 | 10.41 | 0.000 | 14.96392 | 21.9042 |
| UT | 18.5741 | 4.791901 | 3.88 | 0.000 | 9.182129 | 27.96607 |
| VA | 21.17157 | 1.824245 | 11.61 | 0.000 | 17.59611 | 24.74704 |
| VT | 11.1305 | 12.29393 | 0.91 | 0.365 | -12.9652 | 35.2262 |
| WA | 13.80203 | 2.233219 | 6.18 | 0.000 | 9.424991 | 18.17907 |
| WI | 23.59605 | 2.065269 | 11.43 | 0.000 | 19.54819 | 27.64391 |
| WV | 17.70574 | 3.993896 | 4.43 | 0.000 | 9.877837 | 25.53365 |
| WY | 12.65573 | 12.72276 | 0.99 | 0.320 | -12.28048 | 37.59193 |
| meduc_ru_hs | 1.870255 | . 3354285 | 5.58 | 0.000 | 1.212826 | 2.527684 |
| meduc_ru_m s | 1.269429 | . 3632962 | 3.49 | 0.000 | . 5573805 | 1.981478 |
| precare_2t~r | -2.723321 | . 3144676 | -8.66 | 0.000 | -3.339668 | -2.106975 |
| precare_3t~r | -5.439155 | . 5923826 | -9.18 | 0.000 | -6.600206 | -4.278104 |
| precare_none | 31.86282 | . 858685 | 37.11 | 0.000 | 30.17983 | 33.54582 |
| smoked | 1.510182 | . 4679424 | 3.23 | 0.001 | . 5930299 | 2.427334 |
| county_pov~1 | -3.743992 | 3.428558 | -1.09 | 0.275 | -10.46385 | 2.975871 |
| physici~1000 | -. 1200094 | . 0752585 | -1.59 | 0.111 | -. 2675136 | . 0274949 |
| smoked_unk~n | 18.05519 | 1.637276 | 11.03 | 0.000 | 14.84619 | 21.2642 |
| precare_un~n | 10.89554 | . 6234448 | 17.48 | 0.000 | 9.673611 | 12.11747 |
| meduc_ru_u n | 25.99647 | . 9748355 | 26.67 | 0.000 | 24.08583 | 27.90712 |
| _cons | -11.2614 | 1.831551 | -6.15 | 0.000 | -14.85118 | -7.671618 |

reg neonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county_poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mblack==1

## 7.2c Combined factors Model for Neonatal Period (Whites)

. reg neonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_o1der AK $>$ AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC S $>$ D TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county $>$ _poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mwhite==1

| Source | SS | df | MS |  | $\begin{aligned} & \text { Number of obs }=3222929 \\ & F(70,3222858)=96.84 \end{aligned}$ |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Mode 1 | $\begin{aligned} & 25122425.3 \quad 70 \\ & 1.1944 \mathrm{e}+103222858 \end{aligned}$ |  | $\begin{aligned} & 358891.79 \\ & 3706.05637 \end{aligned}$ |  | Prob > F | $=0.0000$ |
| Residual |  |  |  | R -squared | $=0.0021$ |
| Total | $1.1969 \mathrm{e}+103222928$ |  |  | 3713.77078 |  | Root MSE | $=60.877$ |
| neonata~1000 | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Con | Interval] |
| mb1ack | (dropped) |  |  |  |  |  |
| married | -. 6447421 | . 086573 | -7.45 | 0.000 | -. 8144221 | -. 4750621 |
| age_11_16 | 2.552746 | . 3109418 | 8.21 | 0.000 | 1.943311 | 3.162181 |
| age_17_18 | $\begin{array}{r} .741611 \\ .1674833 \end{array}$ | . 1885303 | 3.93 | 0.000 | . 3720982 | 1.111124 |
| age_19_21 |  | . 1231306 | 1.36 | 0.174 | -. 0738483 | . 4088149 |
| age_22_25 | $\begin{array}{r} .1674833 \\ -.0427532 \end{array}$ | . 1010062 | -0.42 | 0.672 | -. 2407218 | . 1552155 |
| age_31_35 | $.2255867$ | . 0973841 | 2.32 | 0.021 | . 0347173 | . 416456 |
| age_36_40 | $\begin{aligned} & 1.392702 \\ & 2.929335 \end{aligned}$ | . 1291511 | 10.78 | 0.000 | 1.13957 | 1.645833 |
| age_41_45 |  | . 2757922 | 10.62 | 0.000 | 2.388792 | 3.469878 |
| age_46_older | $\begin{aligned} & 2.929335 \\ & 6.815697 \end{aligned}$ | 1.214296 | 5.61 | 0.000 | 4.43572 | 9.195675 |
| AK | $\begin{aligned} & 6.815697 \\ & 3.343671 \end{aligned}$ | . 9106971 | 3.67 | 0.000 | 1.558737 | 5.128605 |
| AL | $\begin{aligned} & 5.98502 \\ & 6.30211 \end{aligned}$ | . 5992565 | 9.99 | 0.000 | 4.810499 | 7.159542 |
| AR |  | . 6235815 | 10.11 | 0.000 | 5.079912 | 7.524308 |
| AZ | 5.966983 | . 5527864 | 10.79 | 0.000 | 4.883542 | 7.050425 |
| CO | 6.178926 | . 5716901 | 10.81 | 0.000 | 5.058434 | 7.299419 |
| CT | 5.853354 | . 6146987 | 9.52 | 0.000 | 4.648566 | 7.058142 |
| DC | 6.033167 | 1.377747 | 4.38 | 0.000 | 3.332832 | 8.733502 |
| DE | 6.291042 | . 8563533 | 7.35 | 0.000 | 4.61262 | 7.969464 |
| FL | 5.369123 | . 5403615 | 9.94 | 0.000 | 4.310034 | 6.428213 |
| GA | 5.584965 | . 554377 | 10.07 | 0.000 | 4.498406 | 6.671524 |
| HI | 5.738389 | . 992659 | 5.78 | 0.000 | 3.792812 | 7.683965 |
| IA | 5.393721 | . 6096103 | 8.85 | 0.000 | 4.198907 | 6.588536 |
| ID | 5.890013 | . 6602952 | 8.92 | 0.000 | 4.595858 | 7.184168 |
| IL | 5.784985 | . 5423591 | 10.67 | 0.000 | 4.72198 | 6.847989 |
| IN | 6.52263 | . 5621244 | 11.60 | 0.000 | 5.420886 | 7.624374 |
| KS | 6.591315 | . 6106011 | 10.79 | 0.000 | 5.394559 | 7.788072 |
| KY | 5.472199 | . 5854093 | 9.35 | 0.000 | 4.324817 | 6.61958 |
| LA | 6.565515 | . 6080244 | 10.80 | 0.000 | 5.373809 | 7.757222 |
| MA | 5.60486 | . 5734893 | 9.77 | 0.000 | 4.480842 | 6.728879 |
| MD | 6.224625 | . 5967559 | 10.43 | 0.000 | 5.055004 | 7.394245 |
| ME | 6.279458 | . 7376422 | 8.51 | 0.000 | 4.833705 | 7.725211 |
| MI | 5.702788 | . 5382637 | 10.59 | 0.000 | 4.64781 | 6.757766 |
| MN | 5.019698 | . 5702274 | 8.80 | 0.000 | 3.902073 | 6.137324 |
| MO | 6.055635 | . 5688726 | 10.64 | 0.000 | 4.940665 | 7.170605 |
| MS | 5.635442 | . 6544645 | 8.61 | 0.000 | 4.352715 | 6.91817 |
| MT | 4.426024 | . 8013244 | 5.52 | 0.000 | 2.855456 | 5.996592 |
| NC | 6.362303 | . 5557343 | 11.45 | 0.000 | 5.273083 | 7.451523 |
| ND | 6.788121 | . 8915382 | 7.61 | 0.000 | 5.040737 | 8.535504 |
| NE | 6.400093 | . 6516744 | 9.82 | 0.000 | 5.122834 | 7.677352 |
| NH | 6.144691 | . 7349853 | 8.36 | 0.000 | 4.704146 | 7.585236 |
| NJ | 5.194946 | . 5602307 | 9.27 | 0.000 | 4.096913 | 6.292978 |
| NM | 4.869669 | . 6455922 | 7.54 | 0.000 | 3.604331 | 6.135007 |
| NV | 4.342542 | . 6219884 | 6.98 | 0.000 | 3.123467 | 5.561618 |
| NY | 5.701871 | . 5380972 | 10.60 | 0.000 | 4.64722 | 6.756522 |
| OH | 5.630406 | . 5454122 | 10.32 | 0.000 | 4.561417 | 6.699395 |
| OK | 6.145366 | . 5965457 | 10.30 | 0.000 | 4.976158 | 7.314575 |
| OR | 5.75012 | . 5937946 | 9.68 | 0.000 | 4.586304 | 6.913937 |
| PA | 5.903238 | 5429741 | 10.87 | 0.000 | 4.839028 | 6.967448 |
| RI | 5.626723 | 7738125 | 7.27 | 0.000 | 4.110078 | 7.143369 |
| SC | 6.764768 | 6090287 | 11.11 | 0.000 | 5.571093 | 7.958443 |
| SD | 7.008416 | 8185941 | 8.56 | 0.000 | 5.404 | 8.612831 |
| TN | 5.382902 | . 5732402 | 9.39 | 0.000 | 4.259371 | 6.506432 |
| TX | 5.277502 | 5283198 | 9.99 | 0.000 | 4.242013 | 6.31299 |
| UT | 5.515913 | 5848319 | 9.43 | 0.000 | 4.369664 | 6.662163 |
| VA | 5.960537 | . 564998 | 10.55 | 0.000 | 4.853161 | 7.067913 |
| VT | 4.931113 | . 919535 | 5.36 | 0.000 | 3.128856 | 6.733369 |
| WA | 4.252085 | 5650798 | 7.52 | 0.000 | 3.144548 | 5.359621 |
| WI | 5.423433 | 5735909 | 9.46 | 0.000 | 4.299215 | 6.547651 |
| WV | 6.542207 | . 672177 | 9.73 | 0.000 | 5.224764 | 7.85965 |
| WY | 8.173584 | . 922307 | 8.86 | 0.000 | 6.365894 | 9.981273 |
| meduc_ru_hs | . 0076158 | . 1019575 | 0.07 | 0.940 | -. 1922173 | . 2074488 |
| meduc_ru_m~s | -. 743592 | . 1034505 | -7.19 | 0.000 | -. 9463513 | -. 5408327 |
| precare_2t~r | -. 6365303 | . 1066952 | -5.97 | 0.000 | -. 8456492 | -. 4274114 |
| precare_3t~r | -1.342096 | . 2198239 | -6.11 | 0.000 | -1.772943 | -. 9112491 |
| precare_none | 17.38848 | . 3653362 | 47.60 | 0.000 | 16.67243 | 18.10453 |
| smoked | 1.002257 | . 1207472 | 8.30 | 0.000 | . 7655968 | 1.238917 |
| county_pov~1 | -1.360181 | 1.068801 | -1.27 | 0.203 | -3.454993 | . 7346316 |
| physici~1000 | -. 099966 | . 0245187 | -4.08 | 0.000 | -. 1480217 | -. 0519103 |
| smoked_unk~n | 4.933653 | . 510193 | 9.67 | 0.000 | 3.933693 | 5.933614 |
| precare_un~n | 6.78384 | . 2226989 | 30.46 | 0.000 | 6.347358 | 7.220322 |
| meduc_ru_u n | $\begin{array}{r} 12.81689 \\ -1.387068 \end{array}$ | . 3189935 | 40.18 | 0.000 | 12.19168 | 13.44211 |
| _cons |  | . 5531649 | -2.51 | 0.012 | -2.471251 | -. 3028838 |

reg neonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county_poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if mwhite==1
7.3a Combined factors Model for Postneonatal Period (Blacks \& Whites)
. reg postneonata1_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_o1der $>$ AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC $>$ SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked count $>$ y_poverty_leve1 physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if neonatal != 1 \& (mwhite==1 mbl > ack==1)

| Source | SS |  | df |
| ---: | :--- | :---: | :---: |
| Mode1 | 9852350.29 | 71 | 138765.497 |
| Residua1 | $\mathbf{8 . 7 7 0 9 e + 0 9 3 8 2 1 4 5 4}$ | 2295.16799 |  |
| Tota1 | $\mathbf{8 . 7 8 0 7 e}+093821525$ | 2297.70347 |  |


| Number of obs | $=3821526$ |
| ---: | :--- | ---: |
| $\mathrm{~F}(71,3821454)$ | $=60.46$ |
| Prob $>\mathrm{F}$ | $=0.0000$ |
| $\mathrm{R}-$ squared | $=0.0011$ |
| Adj R-squared | $=0.0011$ |
| Root MSE | $=47.908$ |


| postneo~1000 |
| ---: |
| mblack |
| married |
| age_11_16 |
| age_17_18 |
| age_19_21 |
| age_22_25 |
| age_31_35 |
| age_36_40 |
| age_41_45 |
| age_46_older |


| AK | 2.192213 2.880726 | . 6829165 | 3.21 6.85 | 0.001 0.000 | .8537204 2.055998 | 3.530705 3.705453 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| AR | 2.399717 | . 4451261 | 5.39 | 0.000 | 1.527286 | 3.272149 |
| AZ | 1.760536 | . 4009765 | 4.39 | 0.000 | . 9746366 | 2.546436 |
| CO | 1.599909 | . 4153783 | 3.85 | 0.000 | . 7857824 | 2.414036 |
| CT | 1.39833 | . 443505 | 3.15 | 0.002 | . 5290759 | 2.267584 |
| DC | 2.064211 | . 6709999 | 3.08 | 0.002 | . 7490748 | 3.379347 |
| DE | 2.501907 | . 5917926 | 4.23 | 0.000 | 1.342014 | 3.661799 |
| FL | 2.05381 | . 3861036 | 5.32 | 0.000 | 1.297061 | 2.81056 |
| GA | 2.230775 | . 392847 | 5.68 | 0.000 | 1.460808 | 3.000741 |
| HI | 1.117197 | . 7344342 | 1.52 | 0.128 | -. 3222684 | 2.556662 |
| IA | 1.98367 | . 4463857 | 4.44 | 0.000 | 1.108769 | 2.85857 |
| ID | 1.790345 | . 4916208 | 3.64 | 0.000 | . 8267859 | 2.753905 |
| IL | 2.176487 | . 388307 | 5.61 | 0.000 | 1.415419 | 2.937555 |
| IN | 2.078543 | . 404971 | 5.13 | 0.000 | 1.284814 | 2.872272 |
| KS | 2.558071 | . 4448503 | 5.75 | 0.000 | 1.68618 | 3.429962 |
| KY | 2.279498 | . 4240739 | 5.38 | 0.000 | 1.448328 | 3.110668 |
| LA | 3.201421 | . 4189335 | 7.64 | 0.000 | 2.380326 | 4.022515 |
| MA | 1.216535 | . 4135251 | 2.94 | 0.003 | . 40604 | 2.027029 |
| MD | 1.896547 | . 4156884 | 4.56 | 0.000 | 1.081813 | 2.711282 |
| ME | 1.649796 | . 5536166 | 2.98 | 0.003 | . 5647269 | 2.734865 |
| MI | 1.851852 | . 3861051 | 4.80 | 0.000 | 1.0951 | 2.608605 |
| MN | 1.401531 | . 4112679 | 3.41 | 0.001 | . 5954601 | 2.207601 |
| MO | 1.907617 | . 408624 | 4.67 | 0.000 | 1.106728 | 2.708506 |
| MS | 2.74306 | . 4409106 | 6.22 | 0.000 | 1.878891 | 3.607229 |
| MT | 2.144511 | . 6073688 | 3.53 | 0.000 | . 9540895 | 3.334932 |
| NC | 2.124799 | . 3963284 | 5.36 | 0.000 | 1.348009 | 2.901589 |
| ND | 1.424994 | . 6788897 | 2.10 | 0.036 | . 0943942 | 2.755594 |
| NE | 1.689253 | . 4783628 | 3.53 | 0.000 | . 7516788 | 2.626827 |
| NH | 1.201222 | . 5506449 | 2.18 | 0.029 | . 1219774 | 2.280467 |
| NJ | 1.419062 | . 4006073 | 3.54 | 0.000 | . 6338855 | 2.204238 |
| NM | 1.901682 | . 4784606 | 3.97 | 0.000 | . 9639166 | 2.839448 |
| NV | 1.37714 | . 4529903 | 3.04 | 0.002 | . 4892951 | 2.264985 |
| NY | 1.514692 | . 3850205 | 3.93 | 0.000 | . 7600654 | 2.269319 |
| OH | 2.265293 | . 3909349 | 5.79 | 0.000 | 1.499075 | 3.031512 |
| OK | 2.430835 | . 431911 | 5.63 | 0.000 | 1.584305 | 3.277366 |
| OR | 1.815727 | . 4352126 | 4.17 | 0.000 | . 9627261 | 2.668729 |
| PA | 1.797051 | . 3874715 | 4.64 | 0.000 | 1.03762 | 2.556481 |
| RI | 1.161632 | . 5679208 | 2.05 | 0.041 | . 0485269 | 2.274736 |
| SC | 2.175606 | . 4235454 | 5.14 | 0.000 | 1.345472 | 3.00574 |
| SD | 1.866827 | . 6195925 | 3.01 | 0.003 | . 6524473 | 3.081206 |
| TN | 2.182412 | . 4089387 | 5.34 | 0.000 | 1.380906 | 2.983917 |
| TX | 1.885841 | . 3791204 | 4.97 | 0.000 | 1.142778 | 2.628904 |
| UT | 1.716643 | . 4281719 | 4.01 | 0.000 | . 8774407 | 2.555844 |
| VA | 2.169175 | . 4026061 | 5.39 | 0.000 | 1.380081 | 2.958268 |
| VT | 1.636072 | . 7020669 | 2.33 | 0.020 | . 2600457 | 3.012098 |
| WA | 1.79237 | . 4094377 | 4.38 | 0.000 | . 9898862 | 2.594853 |
| WI | 2.047415 | . 4141847 | 4.94 | 0.000 | 1.235628 | 2.859202 |
| WV | 1.924341 | . 4985004 | 3.86 | 0.000 | . 9472981 | 2.901384 |
| WY | 2.409579 | . 7052954 | 3.42 | 0.001 | 1.027225 | 3.791933 |
| meduc_ru_hs | -. 4407418 | . 0726555 | -6.07 | 0.000 | -. 5831441 | -. 2983395 |
| meduc_ru_m~s | -. 9257995 | . 0744612 | -12.43 | 0.000 | -1.071741 | -. 7798583 |
| precare_2t~r | . 321455 | . 0745533 | 4.31 | 0.000 | . 1753332 | . 4675768 |
| precare_3t~r | . 3711222 | . 1504035 | 2.47 | 0.014 | . 0763367 | . 6659077 |
| precare_none | 4.246006 | . 2436553 | 17.43 | 0.000 | 3.76845 | 4.723562 |
| smoked | 2.452697 | . 0882642 | 27.79 | 0.000 | 2.279703 | 2.625692 |
| county_pov~1 | . 4698482 | . 7447306 | 0.63 | 0.528 | -. 9897975 | 1.929494 |
| physici~1000 | -. 0385904 | . 0171166 | -2.25 | 0.024 | -. 0721384 | -. 0050425 |
| smoked_unk~n | 1.704495 | . 3648709 | 4.67 | 0.000 | . 989361 | 2.419629 |
| precare_un~n | 1.183181 | . 1542896 | 7.67 | 0.000 | . 880779 | 1.485583 |
| meduc_ru_u~n | -. 0572733 | . 2263344 | -0.25 | 0.800 | -. 5008808 | . 3863342 |
| _cons | . 4069911 | . 3949752 | 1.03 | 0.303 | -. 3671463 | 1.181128 |

reg postneonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county_poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if neonatal != 1 \& (mwhite==1 | mblack==1)

## 7.3b Combined factors Model for Postneonatal Period (Blacks)

. reg postneonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older $>$ AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC $>$ SD TN TX UT VA VT WA WI WV WY meduc ru hs meduc ru moths precare 2trimestr precare 3trimestr precare none smoked count $>$ y_poverty_leve1 physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if neonatal != 1 \& mblack==1

| Source | SS df | MS |
| :---: | :---: | :---: |
| Mode 1 | 2493551.8770 | 35622.1695 |
| Residual | $2.6479 \mathrm{e}+09610540$ | 4336.98247 |
| Tota 1 | $2.6504 \mathrm{e}+09610610$ | 4340.56899 |


| Number of obs | $=$ | 610611 |
| :--- | ---: | ---: |
| F $(70,610540)$ | $=$ | 8.21 |
| Prob $>$ F | $=$ | 0.0000 |
| R-squared | $=0.0009$ |  |
| Adj R-squared | $=0.0008$ |  |
| Root MSE | $=65.856$ |  |

postneo~1000 Coef. Std. Err. $\quad$ t $\mathrm{P}>|\mathrm{t}|$ [95\% Conf. Interval

| mb1ack |
| ---: |
| married |
| age_11_16 |
| 17 |

age_17-18
age_22_25
age_31_35
age 3640
age_36_40
age_41_45
age_46_older
$\qquad$

| AK | 13.40261 | 3.539189 | 3.79 | 0.000 | 6.46591 | 20.3393 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| AL | 5.34998 | 1.321296 | 4.05 | 0.000 | 2.760282 | 7.939678 |
| AR | 4.751822 | 1.445574 | 3.29 | 0.001 | 1.918544 | 7.5851 |
| AZ | 3.057232 | 1.656481 | 1.85 | 0.065 | -. 1894175 | 6.303881 |
| CO | 5.596549 | 1.733843 | 3.23 | 0.001 | 2.198274 | 8.994825 |
| CT | 4.461688 | 1.531366 | 2.91 | 0.004 | 1.460259 | 7.463117 |
| DC | 4.786707 | 1.55777 | 3.07 | 0.002 | 1.733528 | 7.839887 |
| DE | 5.389568 | 1.73588 | 3.10 | 0.002 | 1.987299 | 8.791837 |
| FL | 4.661575 | 1.263165 | 3.69 | 0.000 | 2.185813 | 7.137337 |
| GA | 4.836246 | 1.263656 | 3.83 | 0.000 | 2.35952 | 7.312972 |
| HI | 6.515433 | 3.031434 | 2.15 | 0.032 | . 5739187 | 12.45695 |
| IA | 3.047959 | 2.107448 | 1.45 | 0.148 | -1.082571 | 7.178489 |
| ID | . 4157291 | 6.398787 | 0.06 | 0.948 | -12.12569 | 12.95715 |
| IL | 5.015989 | 1.279191 | 3.92 | 0.000 | 2.508815 | 7.523162 |
| IN | 4.58159 | 1.396919 | 3.28 | 0.001 | 1.843673 | 7.319507 |
| KS | 5.622732 | 1.742922 | 3.23 | 0.001 | 2.206661 | 9.038803 |
| KY | 4.70419 | 1.54803 | 3.04 | 0.002 | 1.670101 | 7.738278 |
| LA | 5.204902 | 1. 300384 | 4.00 | 0.000 | 2.656192 | 7.753612 |
| MA | 2.664683 | 1.431936 | 1.86 | 0.063 | -. 1418652 | 5.471232 |
| MD | 4.208185 | 1.299295 | 3.24 | 0.001 | 1.66161 | 6.754761 |
| ME | . 7101365 | 4.587147 | 0.15 | 0.877 | -8.280524 | 9.700797 |
| MI | 4.682222 | 1.28422 | 3.65 | 0.000 | 2.165192 | 7.199251 |
| MN | 3.193618 | 1.44442 | 2.21 | 0.027 | . 3626013 | 6.024634 |
| MO | 4.627398 | 1.372521 | 3.37 | 0.001 | 1.9373 | 7.317495 |
| MS | 5.369913 | 1.325277 | 4.05 | 0.000 | 2.772413 | 7.967413 |
| MT | -. 0824628 | 9.128649 | -0.01 | 0.993 | -17.97432 | 17.8094 |
| NC | 4.64808 | 1.287 | 3.61 | 0.000 | 2.125601 | 7.17056 |
| ND | 11.29605 | 6.833161 | 1.65 | 0.098 | -2.096728 | 24.68882 |
| NE | 3.932216 | 2.072612 | 1.90 | 0.058 | -. 1300369 | 7.994468 |
| NH | . 8133104 | 4.336295 | 0.19 | 0.851 | -7.685688 | 9.312309 |
| NJ | 3.03962 | 1.313512 | 2.31 | 0.021 | . 4651783 | 5.614062 |
| NM | 3.92412 | 3.184645 | 1.23 | 0.218 | -2.317681 | 10.16592 |
| NV | 4.179688 | 1.717987 | 2.43 | 0.015 | . 8124883 | 7.546887 |
| NY | 3.898096 | 1.268535 | 3.07 | 0.002 | 1.411809 | 6.384383 |
| OH | 4.624211 | 1.296873 | 3.57 | 0.000 | 2.082381 | 7.166041 |
| OK | 4.118406 | 1. 54665 | 2.66 | 0.008 | 1.087021 | 7.149791 |
| OR | 3.721607 | 2.364601 | 1.57 | 0.116 | -. 912935 | 8.356149 |
| PA | 3.471258 | 1.262057 | 2.75 | 0.006 | . 9976673 | 5.944848 |
| RI | 4.024848 | 2.236491 | 1.80 | 0.072 | -. 3586019 | 8.408298 |
| SC | 4.809323 | 1.313903 | 3.66 | 0.000 | 2.234116 | 7.384531 |
| SD | -. 0949993 | 5.586957 | -0.02 | 0.986 | -11.04525 | 10.85526 |
| TN | 3.861652 | 1.329649 | 2.90 | 0.004 | 1.255583 | 6.467722 |
| TX | 4.620413 | 1.266514 | 3.65 | 0.000 | 2.138087 | 7.102739 |
| UT | 7.092082 | 3.392433 | 2.09 | 0.037 | . 4430219 | 13.74114 |
| VA | 4.565962 | 1.304185 | 3.50 | 0.000 | 2.009803 | 7.122122 |
| VT | . 0926458 | 8.662604 | 0.01 | 0.991 | -16.88578 | 17.07107 |
| WA | 2.786087 | 1.58975 | 1.75 | 0.080 | -. 3297715 | 5.901945 |
| WI | 6.87362 | 1.473914 | 4.66 | 0.000 | 3.984797 | 9.762443 |
| WV | 3.769078 | 2.830373 | 1.33 | 0.183 | -1.778363 | 9.316519 |
| WY | . 5733934 | 8.964576 | 0.06 | 0.949 | -16.99689 | 18.14367 |
| meduc_ru_hs | -1.039084 | . 2372469 | -4.38 | 0.000 | -1.504081 | -. 574088 |
| meduc_ru_m~s | -1.714426 | . 2569521 | -6.67 | 0.000 | -2.218044 | -1.210809 |
| precare_2t~r | . 3757579 | . 2222222 | 1.69 | 0.091 | -. 0597906 | . 8113063 |
| precare_3t~r | . 4229389 | . 4180039 | 1.01 | 0.312 | -. 3963354 | 1.242213 |
| precare_none | 5.665819 | . 6173676 | 9.18 | 0.000 | 4.455798 | 6.875839 |
| smoked | 3.950311 | . 331421 | 11.92 | 0.000 | 3.300737 | 4.599886 |
| county_pov~1 | 2.79042 | 2.424817 | 1.15 | 0.250 | -1.962144 | 7.542983 |
| physici~1000 | -. 0571291 | . 0532379 | -1.07 | 0.283 | -. 1614736 | . 0472154 |
| smoked_unk~n | 4.654356 | 1.173476 | 3.97 | 0.000 | 2.354381 | 6.954332 |
| precare_un~n | 2.104397 | . 4434552 | 4.75 | 0.000 | 1.235239 | 2.973555 |
| meduc_ru_u n | -1.487918 | . 6983219 | -2.13 | 0.033 | -2.856607 | -. 11923 |
| _cons | . 1834537 | 1.309114 | 0.14 | 0.889 | -2.382367 | 2.749275 |

reg postneonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county_poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if neonatal != 1 \& mblack==1

## 7.3c Combined factors Model for Postneonatal Period (Whites)

. reg postneonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_o1der $>$ AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC $>$ SD TN TX UT VA VT WA WI WV WY meduc ru hs meduc ru moths precare 2trimestr precare 3trimestr precare none smoked count $>y$ y_poverty_leve1 physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if neonatal != 1 \& mwhite==1

| Source | SS |  | df |
| ---: | :---: | :---: | :---: |
| Mode1 | 4903719.76 | 70 | 70053.1395 |
| Residua1 | $6.1224 \mathrm{e}+093210844$ | 1906.77561 |  |
| Tota1 | $6.1273 \mathrm{e}+093210914$ | 1908.26125 |  |


| Number of obs | $=3210915$ |
| ---: | :--- | ---: |
| $\mathrm{~F}(70,3210844)$ | $=36.74$ |
| Prob $>\mathrm{F}$ | $=0.0000$ |
| R-squared | $=0.0008$ |
| Adj R-squared | $=0.0008$ |
| Root MSE | $=43.667$ |

postneo~1000 Coef. Std. Err
[95\% Conf. Interval]

| mb1ack | (dropped) |  |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| married | -.4351403 | .062227 | -6.99 | 0.000 | -.557103 | -.3131775 |
| age_11_16 | 1.393335 | .223771 | 6.23 | 0.000 | .9547518 | 1.83191 |
| age_17.18 | .6782344 | .1355511 | 5.00 | 0.000 | .412559 | .943909 |


| age_17_18 | .6782344 | .1355511 | 5.00 | 0.000 | .412559 | .943909 |
| ---: | ---: | ---: | ---: | ---: | ---: | ---: |
| age_19_21 | .9384168 | .0884941 | 10.60 | 0.000 | .7649715 | 1.11186 |
| age_22_25 | .3342426 | .0725766 | 4.61 | 0.000 | .1919949 | .476490 |
| age_31_35 | .004814 | .069968 | 0.07 | 0.945 | -.1323207 | .141948 |

age_36_40
age_41-45
age_46_older


| AL | 2.349873 | . 4320263 | 5.44 | 0.000 | 1.503117 | 3.19663 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| AR | 1.87496 | . 4494846 | 4.17 | 0.000 | . 9939856 | 2.755934 |
| AZ | 1.271274 | . 3987245 | 3.19 | 0.001 | . 4897882 | 2.05276 |
| CO | . 9447424 | . 4123013 | 2.29 | 0.022 | . 1366464 | 1.752838 |
| CT | . 7204794 | . 4430666 | 1.63 | 0.104 | -. 1479156 | 1.588874 |
| DC | . 7542496 | . 990573 | 0.76 | 0.446 | -1.187238 | 2.695738 |
| DE | 1.795155 | . 6164692 | 2.91 | 0.004 | . 586897 | 3.003413 |
| FL | 1.455051 | . 3898369 | 3.73 | 0.000 | . 690984 | 2.219117 |
| GA | 1.652079 | . 3998974 | 4.13 | 0.000 | . 8682945 | 2.435864 |
| HI | . 219067 | . 7142407 | 0.31 | 0.759 | -1.180819 | 1.618954 |
| IA | 1.455044 | . 4394138 | 3.31 | 0.001 | . 5938086 | 2.31628 |
| ID | 1.225303 | . 4757601 | 2.58 | 0.010 | . 2928301 | 2.157776 |
| IL | 1.503798 | . 3912753 | 3.84 | 0.000 | . 7369118 | 2.270684 |
| IN | 1.500542 | . 4054421 | 3.70 | 0.000 | . 70589 | 2.295194 |
| KS | 1.922964 | . 4401905 | 4.37 | 0.000 | 1.060206 | 2.785722 |
| KY | 1.753093 | . 4220827 | 4.15 | 0.000 | . 9258261 | 2.580361 |
| LA | 2.940462 | . 4383092 | 6.71 | 0.000 | 2.081392 | 3.799533 |
| MA | . 7736132 | . 4135728 | 1.87 | 0.061 | -. 0369749 | 1.584201 |
| MD | 1.488907 | . 4302543 | 3.46 | 0.001 | . 6456241 | 2.332191 |
| ME | 1.130189 | . 531259 | 2.13 | 0.033 | . 0889399 | 2.171438 |
| MI | 1.17847 | . 3882714 | 3.04 | 0.002 | . 417472 | 1.939469 |
| MN | . 8740654 | . 4111933 | 2.13 | 0.034 | . 0681411 | 1.67999 |
| MO | 1.292151 | . 4102807 | 3.15 | 0.002 | . 4880151 | 2.096286 |
| MS | 2.022971 | . 471597 | 4.29 | 0.000 | 1.098657 | 2.947284 |
| MT | 1.606033 | . 576621 | 2.79 | 0.005 | . 4758763 | 2.73619 |
| NC | 1.550232 | . 4008567 | 3.87 | 0.000 | . 7645675 | 2.335897 |
| ND | . 7312454 | . 6417523 | 1.14 | 0.255 | -. 5265665 | 1.989057 |
| NE | 1.111319 | . 4696087 | 2.37 | 0.018 | . 1909028 | 2.031736 |
| NH | . 6489575 | . 5293665 | 1.23 | 0.220 | -. 3885821 | 1.686497 |
| NJ | 1.027367 | . 4040501 | 2.54 | 0.011 | . 2354427 | 1.819291 |
| NM | 1.3841 | . 465245 | 2.97 | 0.003 | . 4722366 | 2.295964 |
| NV | . 797641 | . 4482618 | 1.78 | 0.075 | -. 0809363 | 1.676218 |
| NY | . 9471178 | . 3882155 | 2.44 | 0.015 | . 1862292 | 1.708006 |
| OH | 1.704179 | . 3934646 | 4.33 | 0.000 | . 9330024 | 2.475356 |
| OK | 1.967201 | . 4300846 | 4.57 | 0.000 | 1.124251 | 2.810152 |
| OR | 1.265147 | . 4281364 | 2.96 | 0.003 | . 4260152 | 2.10428 |
| PA | 1.317574 | . 3917603 | 3.36 | 0.001 | . 5497379 | 2.08541 |
| RI | . 5289326 | . 5572098 | 0.95 | 0.342 | -. 5631789 | 1.621044 |
| SC | 1.52376 | . 4390809 | 3.47 | 0.001 | . 6631766 | 2.384343 |
| SD | 1.366604 | . 5894411 | 2.32 | 0.020 | . 2113206 | 2.521888 |
| TN | 1.806298 | . 413389 | 4.37 | 0.000 | . 9960702 | 2.616526 |
| TX | 1.320292 | . 3812127 | 3.46 | 0.001 | . 5731288 | 2.067456 |
| UT | 1.081189 | . 4216435 | 2.56 | 0.010 | . 2547829 | 1.907596 |
| VA | 1.644531 | . 4074879 | 4.04 | 0.000 | . 8458692 | 2.443193 |
| VT | 1.090921 | . 6615113 | 1.65 | 0.099 | -. 2056175 | 2.38746 |
| WA | 1.328199 | . 4075184 | 3.26 | 0.001 | . 5294769 | 2.12692 |
| WI | 1.176801 | . 4136164 | 2.85 | 0.004 | . 3661279 | 1.987475 |
| WV | 1.421337 | . 4843925 | 2.93 | 0.003 | . 4719445 | 2.370729 |
| WY | 1.867204 | . 6642694 | 2.81 | 0.005 | . 5652592 | 3.169148 |
| meduc_ru_hs | -. 3082298 | . 073283 | -4.21 | 0.000 | -. 4518618 | -. 1645978 |
| meduc_ru_m~s | -. 7786539 | . 0743478 | -10.47 | 0.000 | -. 9243729 | -. 6329349 |
| precare_2t~r | . 3277749 | . 0766578 | 4.28 | 0.000 | . 1775282 | . 4780215 |
| precare_3t~r | . 384133 | . 1578955 | 2.43 | 0.015 | . 0746633 | . 6936027 |
| precare_none | 3.570495 | . 2649062 | 13.48 | 0.000 | 3.051288 | 4.089702 |
| smoked | 2.196159 | . 0868222 | 25.29 | 0.000 | 2.025991 | 2.366328 |
| county_pov~1 | -. 1851909 | . 7679308 | -0.24 | 0.809 | -1.690308 | 1.319926 |
| physici~1000 | -. 0347096 | . 0176189 | -1.97 | 0.049 | -. 0692419 | -. 0001772 |
| smoked_unk~n | 1.10589 | . 3682472 | 3.00 | 0.003 | . 3841387 | 1.827642 |
| precare_un~n | . 9478016 | . 1605708 | 5.90 | 0.000 | . 6330885 | 1.262515 |
| meduc_ru_u n | . 2509262 | . 2307073 | 1.09 | 0.277 | -. 2012519 | . 7031043 |
| _cons | . 8887111 | . 3990274 | 2.23 | 0.026 | . 1066315 | 1.670791 |

reg postneonatal_1000 mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths precare_2trimestr precare_3trimestr precare_none smoked county_poverty_level physicians_per_1000 smoked_unknown precare_unknown meduc_ru_unknown if neonatal != 1 \& mwhite==1

Appendix B. Factors associated with starting prenatal care in first trimester
Prenatal care from the $1^{\text {st }}$ trimester $=a+b *$ Mother is Black $+P_{i}$ Married $+E_{i}$ Education Level $+D_{i}$ Birthday Sept-Dec + Ki State $+L_{i}$ Age Group $+S_{i}$ Smoked $+u$ (Equation 7.1)

Table 4a. Prenatal care started in $1^{\text {st }}$ trimester

| Section 4. | Blacks and Whites |  | Blacks Only |  | Whites Only |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Coefficient | SE | Coefficient | SE | Coefficient | SE |
| mblack | -.067 | .00058 |  |  |  |  |
| BD Sept-to-Dec | .00166 | .0004 | .0094 | .0012 | .00021 | .00044 |
| married | .082 | .0005 | .075 | .0014 | .084 | .0004 |
| high school | .093 | .0006 | .06 | .0016 | .097 | .0006 |
| more than HS | .154 | .0006 | .13 | .0017 | .155 | .00062 |
| smoked | -.035 | .0007 | -.084 | .0022 | -.028 | .0007 |
| smoked <br> unknown | -.1926 | .0029 | -.164 | .0076 | -.194 | .003 |
| educ unknown | -.075 | .0018 | -.134 | .0045 | -.061 | .0019 |
| cons | .93 | .003 | .895 | .008 | .92 | .003 |

$R^{2}=0.0852$ (This highest $R^{2}$ in the paper. Rests of the models have $R^{2}$ less than 0.005 . Reason is, it's almost impossible to predict particular death based socioeconomic variables. From other side whether or not mother will start prenatal in the $1^{\text {st }}$ trimester is more predictable).

Results showed that being married and having higher level of education associated with early start in prenatal care, while smoking during the pregnancy associated with postponing prenatal care.

When baby was born between September and December (variable BD Sep-to-Dec), mother more likely to start prenatal care from the first trimester. Idea behind it is that health insurance detectable resets every new calendar year. Mother who knows that baby is going to be born same calendar year as she got pregnant (born between September to December) will more likely to use prenatal care from the first trimester. Due to large hospital bill at time of the delivery, she will likely to meet her annual deductible anyway. Thus starting prenatal care early would not increase her annual out-of-pocket expenditures to more than co-insurance premium (usually around $20 \%$ of the insurer adjusted medical bill). Effect is small in magnitude, but statistically significant among black mothers. Among white mothers effect is not statistically significant.

Based on results one could speculate that black mothers are more likely to start prenatal care based on annual health insurance payments schedule than white mothers. This follows from budget constraint analysis. On average blacks optimize around smaller budget leading them to choose bundle with less prenatal care. When part of prenatal care cost can be avoided blacks more likely to do so compare to whites.

Table 4.1a Prenatal care started in the 1st trimester (Blacks \& Whites)
. reg precare_1trimestr bd_sept_to_dec mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41 $>$ _ 45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY O $>\bar{H}$ OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths smoked meduc_ru_unknown smoked_unknown if mwh > ite==1 | mblack==1

| Source | SS | $d f$ | MS |
| ---: | :---: | :---: | :---: |
| Mode1 | 57621.011 | 66 | 873.045621 |
| Residua1 | 559406.4663838936 | .145719144 |  |
| Tota1 | 617027.4773839002 | .160726011 |  |

$$
\begin{aligned}
\text { Number of obs } & =3839003 \\
\text { F }(66,3838936) & =5991.29 \\
\text { Prob }>\text { F } & =0.0000 \\
\text { R-squared } & =0.0934 \\
\text { Adj R-squared } & =0.0934 \\
\text { Root MSE } & =.38173
\end{aligned}
$$

| precare_1t~r | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Conf. | Interval] |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| bd_sept_to~c | . 0016639 | . 0004124 | 4.03 | 0.000 | . 0008556 | . 0024723 |
| mblack | -. 0669192 | . 0005798 | -115.41 | 0.000 | -. 0680556 | -. 0657828 |
| married | . 0820989 | . 0004908 | 167.29 | 0.000 | . 081137 | . 0830607 |
| age_11_16 | -. 089194 | . 0016282 | -54.78 | 0.000 | -. 0923853 | -. 0860027 |
| age_17_18 | -. 0379766 | . 0010326 | -36.78 | 0.000 | -. 0400004 | -. 0359527 |
| age_19_21 | -. 039833 | . 0006908 | -57.66 | 0.000 | -. 0411869 | -. 0384792 |
| age_22_25 | -. 0205679 | . 0005792 | -35.51 | 0.000 | -. 0217031 | -. 0194327 |
| age_31_35 | . 005337 | . 000572 | 9.33 | 0.000 | . 004216 | . 0064581 |
| age_36_40 | -. 0051229 | . 0007572 | -6.77 | 0.000 | -. 006607 | -. 0036387 |
| age_41_45 | -. 0355094 | . 0016106 | -22.05 | 0.000 | -. 0386662 | -. 0323526 |
| age_46_older | -. 0612401 | . 0071669 | -8.54 | 0.000 | -. 0752869 | -. 0471933 |
| AK | -. 2742345 | . 0054099 | -50.69 | 0.000 | -. 2848376 | -. 2636314 |
| AL | -. 2020857 | . 0033169 | -60.93 | 0.000 | -. 2085868 | -. 1955846 |
| AR | -. 2308556 | . 0035114 | -65.74 | 0.000 | -. 2377379 | -. 2239733 |
| AZ | -. 2756254 | . 0031607 | -87.20 | 0.000 | -. 2818203 | -. 2694306 |
| CO | -. 2805487 | . 0032703 | -85.79 | 0.000 | -. 2869584 | -. 2741391 |
| CT | -. 2166792 | . 0034881 | -62.12 | 0.000 | -. 2235158 | -. 2098427 |
| DC | -. 2986589 | . 005258 | -56.80 | 0.000 | -. 3089644 | -. 2883534 |
| DE | -. 1929006 | . 004679 | -41.23 | 0.000 | -. 2020714 | -. 1837299 |
| FL | -. 3289022 | . 003035 | -108.37 | 0.000 | -. 3348506 | -. 3229537 |
| GA | -. 207202 | . 003095 | -66.95 | 0.000 | -. 2132681 | -. 201136 |
| HI | -. 2554456 | . 005816 | -43.92 | 0.000 | -. 2668448 | -. 2440463 |
| IA | -. 1969605 | . 0035215 | -55.93 | 0.000 | -. 2038625 | -. 1900584 |
| ID | -. 3698459 | . 0038831 | -95.24 | 0.000 | -. 3774566 | -. 3622351 |
| IL | -. 2444025 | . 0030535 | -80.04 | 0.000 | -. 2503872 | -. 2384178 |
| IN | -. 2532056 | . 0031899 | -79.38 | 0.000 | -. 2594576 | -. 2469536 |
| KS | -. 2096712 | . 0035099 | -59.74 | 0.000 | -. 2165505 | -. 2027919 |
| KY | -. 3161083 | . 0033404 | -94.63 | 0.000 | -. 3226554 | -. 3095613 |
| LA | -. 1667016 | . 0032856 | -50.74 | 0.000 | -. 1731413 | -. 1602619 |
| MA | -. 2075185 | . 0032385 | -64.08 | 0.000 | -. 2138658 | -. 2011712 |
| MD | -. 2467809 | . 0032595 | -75.71 | 0.000 | -. 2531695 | -. 2403924 |
| ME | -. 196958 | . 004381 | -44.96 | 0.000 | -. 2055447 | -. 1883714 |
| MI | -. 2224209 | . 0030405 | -73.15 | 0.000 | -. 2283803 | -. 2164616 |
| MN | -. 2421055 | . 0032261 | -75.05 | 0.000 | -. 2484286 | -. 2357825 |
| MO | -. 1894518 | . 0032213 | -58.81 | 0.000 | -. 1957655 | -. 1831381 |
| MS | -. 1952675 | . 0034624 | -56.40 | 0.000 | -. 2020537 | -. 1884813 |
| MT | -. 226525 | . 0048149 | -47.05 | 0.000 | -. 2359621 | -. 217088 |
| NC | -. 2125564 | . 0031241 | -68.04 | 0.000 | -. 2186796 | -. 2064333 |
| ND | -. 2293822 | . 0053798 | -42.64 | 0.000 | -. 2399265 | -. 2188379 |
| NE | -. 2504778 | . 0037759 | -66.34 | 0.000 | -. 2578785 | -. 2430771 |
| NH | -. 2251004 | . 0043392 | -51.88 | 0.000 | -. 2336051 | -. 2165956 |
| NJ | -. 2919082 | . 0031439 | -92.85 | 0.000 | -. 2980701 | -. 2857463 |
| NM | -. 3674819 | . 0037728 | -97.40 | 0.000 | -. 3748764 | -. 3600874 |
| NV | -. 3435477 | . 0035676 | -96.30 | 0.000 | -. 35054 | -. 3365553 |
| NY | -. 2898876 | . 0030246 | -95.84 | 0.000 | -. 2958157 | -. 2839596 |
| OH | -. 2140951 | . 0030755 | -69.61 | 0.000 | -. 220123 | -. 2080673 |
| OK | -. 2736139 | . 0034068 | -80.31 | 0.000 | -. 2802911 | -. 2669367 |
| OR | -. 260748 | . 0034356 | -75.90 | 0.000 | -. 2674817 | -. 2540143 |
| PA | -. 3684048 | . 0030417 | -121.12 | 0.000 | -. 3743663 | -. 3624432 |
| RI | -. 2103885 | . 0044906 | -46.85 | 0.000 | -. 2191899 | -. 2015872 |
| SC | -. 3496949 | . 0033392 | -104.72 | 0.000 | -. 3562397 | -. 3431502 |
| SD | -. 2675057 | . 0049082 | -54.50 | 0.000 | -. 2771257 | -. 2578857 |
| TN | -. 4072019 | . 0032121 | -126.77 | 0.000 | -. 4134975 | -. 4009063 |
| TX | -. 233834 | . 0029792 | -78.49 | 0.000 | -. 2396731 | -. 2279949 |
| UT | -. 3038409 | . 0033732 | -90.08 | 0.000 | -. 3104521 | -. 2972296 |
| VA | -. 2087501 | . 0031659 | -65.94 | 0.000 | -. 2149552 | -. 202545 |
| VT | -. 208605 | . 0055575 | -37.54 | 0.000 | -. 2194976 | -. 1977124 |
| WA | -. 4602977 | . 0032077 | -143.50 | 0.000 | -. 4665846 | -. 4540107 |
| WI | -. 2161511 | . 0032654 | -66.19 | 0.000 | -. 2225511 | -. 2097511 |
| WV | -. 2155737 | . 0039369 | -54.76 | 0.000 | -. 2232898 | -. 2078576 |
| WY | -. 2162598 | . 0055855 | -38.72 | 0.000 | -. 2272072 | -. 2053124 |
|  | . 0928973 | . 0005748 | 161.61 | 0.000 | . 0917706 | . 0940239 |
| meduc_ru_m~s | . 1538754 | . 0005848 | 263.14 | 0.000 | . 1527293 | . 1550215 |
| smoked | -. 0346571 | . 0006996 | -49.54 | 0.000 | -. 0360282 | -. 033286 |
| meduc_ru_u n | -. 0750901 | . 001777 | -42.26 | 0.000 | -. 0785729 | -. 0716073 |
| smoked_unk~n | -. 1925879 | . 0028713 | -67.07 | 0.000 | -. 1982155 | -. 1869602 |
| _cons | . 928371 | . 0029842 | 311.09 | 0.000 | . 9225221 | . 93422 |

reg precare_1trimestr bd_sept_to_dec mblack married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths smoked meduc_ru_unknown smoked_unknown if mwhite==1 \| mblack==1

Table 4.1b Prenatal care started in the 1st trimester (Blacks)
. reg precare_1trimestr mblack bd_sept_to_dec married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_ $>45$ age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH $>$ OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths smoked meduc_ru_unknown smoked_unknown if mbla $>\mathrm{ck}==1$

| Source | SS | df | MS |
| ---: | ---: | ---: | ---: |
| Mode1 | 10839.576 | 65 | 166.762708 |
| Residua1 | 118177.389616008 | .191843919 |  |
| Tota1 | 129016.965616073 | .209418308 |  |


| Number of obs | $=616074$ |
| ---: | :--- | ---: |
| $\mathrm{~F}(65,616008)$ | $=869.26$ |
| Prob $>$ F | $=0.0000$ |
| R-squared | $=0.0840$ |
| Adj R-squared | $=0.0839$ |
| Root MSE | $=.438$ |


| precare_1t~r | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Conf. | Interval] |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| mb1ack | (dropped) |  |  |  |  |  |
| bd_sept_to~c | . 0094052 | . 0011713 | 8.03 | 0.000 | . 0071095 | . 0117008 |
| married | . 0752653 | . 001378 | 54.62 | 0.000 | . 0725645 | . 0779662 |
| age_11_16 | -. 1456425 | . 0035252 | -41.31 | 0.000 | -. 1525519 | -. 1387332 |
| age_17_18 | -. 0550548 | . 0025032 | -21.99 | 0.000 | -. 0599609 | -. 0501487 |
| age_19_21 | -. 0402704 | . 0018199 | -22.13 | 0.000 | -. 0438373 | -. 0367035 |
| age_22_25 | -. 0154145 | . 0016603 | -9.28 | 0.000 | -. 0186686 | -. 0121603 |
| age_31_35 | -. 0036583 | . 0018961 | -1.93 | 0.054 | -. 0073746 | . 000058 |
| age_36_40 | -. 0129453 | . 0024817 | -5.22 | 0.000 | -. 0178094 | -. 0080811 |
| age_41_45 | -. 0503809 | . 0050925 | -9.89 | 0.000 | -. 060362 | -. 0403997 |
| age_46_older | -. 0657231 | . 0243686 | -2.70 | 0.007 | -. 1134848 | -. 0179615 |
| AK | -. 2235084 | . 0233674 | -9.56 | 0.000 | -. 2693076 | -. 1777091 |
| AL | -. 1925775 | . 0086356 | -22.30 | 0.000 | -. 209503 | -. 175652 |
| AR | -. 2184457 | . 0094674 | -23.07 | 0.000 | -. 2370015 | -. 19989 |
| AZ | -. 213394 | . 0108934 | -19.59 | 0.000 | -. 2347447 | -. 1920432 |
| CO | -. 2792482 | . 0113748 | -24.55 | 0.000 | -. 3015424 | -. 2569539 |
| CT | -. 2146409 | . 0100078 | -21.45 | 0.000 | -. 2342559 | -. 1950258 |
| DC | -. 3179029 | . 0100206 | -31.72 | 0.000 | -. 3375431 | -. 2982628 |
| DE | -. 1433528 | . 0113903 | -12.59 | 0.000 | -. 1656775 | -. 1210282 |
| FL | -. 3518243 | . 0082346 | -42.73 | 0.000 | -. 3679638 | -. 3356849 |
| GA | -. 1912041 | . 0082541 | -23.16 | 0.000 | -. 2073819 | -. 1750263 |
| HI | -. 1667616 | . 0199774 | -8.35 | 0.000 | -. 2059167 | -. 1276065 |
| IA | -. 1762926 | . 0139077 | -12.68 | 0.000 | -. 2035512 | -. 1490341 |
| ID | -. 3169356 | . 0425259 | -7.45 | 0.000 | -. 4002849 | -. 2335863 |
| IL | -. 2435676 | . 0083491 | -29.17 | 0.000 | -. 2599315 | -. 2272036 |
| IN | -. 2688446 | . 0091428 | -29.41 | 0.000 | -. 2867642 | -. 2509251 |
| KS | -. 1810058 | . 0114603 | -15.79 | 0.000 | -. 2034677 | -. 158544 |
| KY | -. 2865763 | . 0101448 | -28.25 | 0.000 | -. 3064598 | -. 2666927 |
| LA | -. 1767081 | . 0084281 | -20.97 | 0.000 | -. 193227 | -. 1601892 |
| MA | -. 196138 | . 0092711 | -21.16 | 0.000 | -. 2143091 | -. 177967 |
| MD | -. 2679311 | . 008453 | -31.70 | 0.000 | -. 2844986 | -. 2513636 |
| ME | -. 1794799 | . 030274 | -5.93 | 0.000 | -. 238816 | -. 1201439 |
| MI | -. 2559736 | . 0083865 | -30.52 | 0.000 | -. 2724107 | -. 2395364 |
| MN | -. 2989172 | . 0094213 | -31.73 | 0.000 | -. 3173827 | -. 2804517 |
| MO | -. 1641335 | . 0089655 | -18.31 | 0.000 | -. 1817056 | -. 1465614 |
| MS | -. 1898658 | . 0086142 | -22.04 | 0.000 | -. 2067494 | -. 1729821 |
| MT | -. 1013023 | . 0606943 | -1.67 | 0.095 | -. 2202611 | . 0176566 |
| NC | -. 2062308 | . 0084108 | -24.52 | 0.000 | -. 2227158 | -. 1897459 |
| ND | -. 1896651 | . 0454179 | -4.18 | 0.000 | -. 2786828 | -. 1006475 |
| NE | -. 2274574 | . 0136329 | -16.68 | 0.000 | -. 2541775 | -. 2007374 |
| NH | -. 27026 | . 0285139 | -9.48 | 0.000 | -. 3261463 | -. 2143738 |
| NJ | -. 3393583 | . 0085653 | -39.62 | 0.000 | -. 356146 | -. 3225705 |
| NM | -. 3461854 | . 0210802 | -16.42 | 0.000 | -. 3875018 | -. 3048689 |
| NV | -. 325292 | . 0112695 | -28.86 | 0.000 | -. 3473798 | -. 3032041 |
| NY | -. 2850762 | . 0082372 | -34.61 | 0.000 | -. 3012209 | -. 2689315 |
| OH | -. 2279127 | . 0084457 | -26.99 | 0.000 | -. 244466 | -. 2113595 |
| OK | -. 2551048 | . 0101453 | -25.15 | 0.000 | -. 2749893 | -. 2352203 |
| OR | -. 2229853 | . 0156096 | -14.29 | 0.000 | -. 2535795 | -. 1923911 |
| PA | -. 4593493 | . 0081782 | -56.17 | 0.000 | -. 4753783 | -. 4433203 |
| RI | -. 2316703 | . 0147576 | -15.70 | 0.000 | -. 2605947 | -. 2027459 |
| SC | -. 365949 | . 0085849 | -42.63 | 0.000 | -. 3827752 | -. 3491229 |
| SD | -. 304671 | . 0370047 | -8.23 | 0.000 | -. 3771991 | -. 232143 |
| TN | -. 5150146 | . 0086282 | -59.69 | 0.000 | -. 5319256 | -. 4981037 |
| TX | -. 1979576 | . 0082684 | -23.94 | 0.000 | -. 2141633 | -. 1817519 |
| UT | -. 3803207 | . 0224316 | -16.95 | 0.000 | -. 4242858 | -. 3363556 |
| VA | -. 1861868 | . 0085081 | -21.88 | 0.000 | -. 2028624 | -. 1695113 |
| VT | -. 2939552 | . 057583 | -5.10 | 0.000 | -. 4068162 | -. 1810943 |
| WA | -. 4856633 | . 0103747 | -46.81 | 0.000 | -. 5059974 | -. 4653292 |
| WI | -. 1576621 | . 0096397 | -16.36 | 0.000 | -. 1765557 | -. 1387686 |
| WV | -. 2321348 | . 0186938 | -12.42 | 0.000 | -. 268774 | -. 1954956 |
| WY | -. 1576646 | . 0596 | -2.65 | 0.008 | -. 2744786 | -. 0408506 |
| meduc_ru_hs | . 0613334 | . 0015685 | 39.10 | 0.000 | . 0582593 | . 0644076 |
| meduc_ru_m~s | . 1301577 | . 0016903 | 77.00 | 0.000 | . 1268448 | . 1334705 |
| smoked | -. 0841665 | . 0021854 | -38.51 | 0.000 | -. 0884499 | -. 0798831 |
| meduc_ru_u~n | -. 1340713 | . 0045311 | -29.59 | 0.000 | -. 1429522 | -. 1251905 |
| smoked_unk~n | -. 1635592 | . 0076501 | -21.38 | 0.000 | -. 1785531 | -. 1485653 |
| _cons | . 8947533 | . 0082268 | 108.76 | 0.000 | . 878629 | . 9108777 |

reg precare_1trimestr mblack bd_sept_to_dec married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths smoked meduc_ru_unknown smoked_unknown if mblack==1

Table 4.1c Prenatal care started in the 1st trimester (Whites)

- reg precare_1trimestr mblack bd_sept_to_dec married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_
$>45$ age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH $>45$ age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH
$>$ OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths smoked meduc_ru_unknown smoked_unknown if mwhi $>$ te==1

| Source | SS | df | MS |
| ---: | :--- | :---: | :---: |
| Mode1 | 40968.9834 | 65 | 630.292053 |
| Residua1 | 440088.6693222863 | .136552087 |  |
| Tota1 | 481057.6533222928 | .149261061 |  |


| Number of obs | $=3222929$ |
| ---: | :--- |
| $\mathrm{~F}(65,3222863)$ | $=4615.76$ |
| Prob $>\mathrm{F}$ | $=0.0000$ |
| R -squared | $=0.0852$ |
| Adj R-squared | $=0.0851$ |
| Root MSE | $=.36953$ |


| precare_1t~r | Coef. | Std. Err. | t | $P>\|t\|$ | [95\% Con | Interval] |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| mblack | (dropped) |  |  |  |  |  |
| bd_sept_to~c | . 0002059 | . 0004365 | 0.47 | 0.637 | -. 0006496 | . 0010613 |
| married | . 0838547 | . 0005227 | 160.44 | 0.000 | . 0828303 | . 0848791 |
| age_11_16 | -. 0714841 | . 0018868 | -37.89 | 0.000 | -. 0751822 | -. 0677861 |
| age_17_18 | -. 0353937 | . 0011439 | -30.94 | 0.000 | -. 0376358 | -. 0331516 |
| age_19_21 | -. 0410578 | . 0007468 | -54.98 | 0.000 | -. 0425214 | -. 0395942 |
| age_22_25 | -. 0223242 | . 0006128 | -36.43 | 0.000 | -. 0235252 | -. 0211231 |
| age_31_35 | . 0065093 | . 0005905 | 11.02 | 0.000 | . 0053519 | . 0076667 |
| age_36_40 | -. 004031 | . 0007829 | -5.15 | 0.000 | -. 0055655 | -. 0024966 |
| age_41_45 | -. 0330975 | . 0016734 | -19.78 | 0.000 | -. 0363773 | -. 0298177 |
| age_46_older | -. 0603119 | . 0073703 | -8.18 | 0.000 | -. 0747574 | -. 0458663 |
| AK | -. 2782565 | . 0055119 | -50.48 | 0.000 | -. 2890596 | -. 2674535 |
| AL | -. 2093625 | . 003622 | -57.80 | 0.000 | -. 2164615 | -. 2022636 |
| AR | -. 2350255 | . 0037704 | -62.33 | 0.000 | -. 2424153 | -. 2276357 |
| AZ | -. 2777305 | . 0033422 | -83.10 | 0.000 | -. 284281 | -. 2711799 |
| CO | -. 2804549 | . 0034505 | -81.28 | 0.000 | -. 2872178 | -. 273692 |
| CT | -. 2174411 | . 0037044 | -58.70 | 0.000 | -. 2247016 | -. 2101805 |
| DC | -. 2619568 | . 0083167 | -31.50 | 0.000 | -. 2782572 | -. 2456563 |
| DE | -. 211793 | . 005184 | -40.86 | 0.000 | -. 2219535 | -. 2016326 |
| FL | -. 3232932 | . 0032599 | -99.17 | 0.000 | -. 3296825 | -. 316904 |
| GA | -. 2176823 | . 0033515 | -64.95 | 0.000 | -. 2242512 | -. 2111135 |
| HI | -. 265199 | . 0060094 | -44.13 | 0.000 | -. 2769772 | -. 2534208 |
| IA | -. 1977956 | . 0036835 | -53.70 | 0.000 | -. 205015 | -. 1905761 |
| ID | -. 3700605 | . 0039919 | -92.70 | 0.000 | -. 3778844 | -. 3622366 |
| IL | -. 2451357 | . 0032715 | -74.93 | 0.000 | -. 2515478 | -. 2387237 |
| IN | -. 2517215 | . 0033951 | -74.14 | 0.000 | -. 2583757 | -. 2450672 |
| KS | -. 2118454 | . 0036915 | -57.39 | 0.000 | -. 2190806 | -. 2046102 |
| KY | -. 3198624 | . 003535 | -90.48 | 0.000 | -. 3267908 | -. 3129339 |
| LA | -. 1653181 | . 0036601 | -45.17 | 0.000 | -. 1724918 | -. 1581444 |
| MA | -. 2096183 | . 0034455 | -60.84 | 0.000 | -. 2163714 | -. 2028652 |
| MD | -. 2360713 | . 003586 | -65.83 | 0.000 | -. 2430996 | -. 2290429 |
| ME | -. 1976515 | . 0044665 | -44.25 | 0.000 | -. 2064056 | -. 1888974 |
| MI | -. 2157513 | . 0032513 | -66.36 | 0.000 | -. 2221238 | -. 2093789 |
| MN | -. 2359184 | . 003428 | -68.82 | 0.000 | -. 2426372 | -. 2291995 |
| MO | -. 194642 | . 0034389 | -56.60 | 0.000 | -. 2013821 | -. 1879018 |
| MS | -. 2057319 | . 0039406 | -52.21 | 0.000 | -. 2134555 | -. 1980084 |
| MT | -. 2274357 | . 0048559 | -46.84 | 0.000 | -. 2369531 | -. 2179183 |
| NC | -. 215433 | . 0033614 | -64.09 | 0.000 | -. 2220213 | -. 2088447 |
| ND | -. 2298625 | . 0054011 | -42.56 | 0.000 | -. 2404485 | -. 2192765 |
| NE | -. 2519583 | . 0039389 | -63.97 | 0.000 | -. 2596784 | -. 2442382 |
| NH | -. 224644 | . 0044281 | -50.73 | 0.000 | -. 233323 | -. 2159651 |
| NJ | -. 2811076 | . 0033701 | -83.41 | 0.000 | -. 2877129 | -. 2745024 |
| NM | -. 367772 | . 0038981 | -94.35 | 0.000 | -. 3754122 | -. 3601318 |
| NV | -. 3453822 | . 0037515 | -92.06 | 0.000 | -. 3527351 | -. 3380294 |
| NY | -. 2927057 | . 0032441 | -90.23 | 0.000 | -. 2990641 | -. 2863473 |
| OH | -. 2121214 | . 0032926 | -64.42 | 0.000 | -. 2185748 | -. 205668 |
| OK | -. 2761225 | . 0036068 | -76.56 | 0.000 | -. 2831918 | -. 2690533 |
| OR | -. 2616017 | . 0035929 | -72.81 | 0.000 | -. 2686436 | -. 2545597 |
| PA | -. 3513702 | . 0032709 | -107.42 | 0.000 | -. 3577809 | -. 3449594 |
| RI | -. 2085521 | . 0046811 | -44.55 | 0.000 | -. 2177269 | -. 1993773 |
| SC | -. 3447534 | . 0036829 | -93.61 | 0.000 | -. 3519717 | -. 3375351 |
| SD | -. 2671304 | . 0049604 | -53.85 | 0.000 | -. 2768526 | -. 2574081 |
| TN | -. 3786935 | . 0034573 | -109.53 | 0.000 | -. 3854698 | -. 3719173 |
| TX | -. 2384998 | . 0031862 | -74.85 | 0.000 | -. 2447446 | -. 2322551 |
| UT | -. 3028241 | . 0035294 | -85.80 | 0.000 | -. 3097416 | -. 2959066 |
| VA | -. 2169266 | . 0034078 | -63.66 | 0.000 | -. 2236057 | -. 2102475 |
| VT | -. 208341 | . 0055594 | -37.48 | 0.000 | -. 2192372 | -. 1974448 |
| WA | -. 4588093 | . 0033935 | -135.20 | 0.000 | -. 4654604 | -. 4521581 |
| WI | -. 223051 | . 0034653 | -64.37 | 0.000 | -. 229843 | -. 2162591 |
| WV | -. 2157721 | . 0040656 | -53.07 | 0.000 | -. 2237404 | -. 2078037 |
| WY | -. 2170718 | . 0055866 | -38.86 | 0.000 | -. 2280213 | -. 2061223 |
| meduc_ru_hs | . 0968942 | . 0006158 | 157.35 | 0.000 | . 0956873 | . 0981011 |
| meduc_ru_m~s | . 1551732 | . 00062 | 250.28 | 0.000 | . 153958 | . 1563884 |
| smoked | -. 0282402 | . 0007312 | -38.62 | 0.000 | -. 0296732 | -. 0268071 |
| meduc_ru_u n | -. 0614674 | . 0019286 | -31.87 | 0.000 | -. 0652474 | -. 0576874 |
| smoked_unk~n | -. 1943333 | . 0030831 | -63.03 | 0.000 | -. 2003762 | -. 1882904 |
| _cons | . 924652 | . 003194 | 289.50 | 0.000 | . 918392 | . 9309121 |

reg precare_1trimestr mblack bd_sept_to_dec married age_11_16 age_17_18 age_19_21 age_22_25 age_31_35 age_36_40 age_41_45 age_46_older AK AL AR AZ CO CT DC DE FL GA HI IA ID IL IN KS KY LA MA MD ME MI MN MO MS MT NC ND NE NH NJ NM NV NY OH OK OR PA RI SC SD TN TX UT VA VT WA WI WV WY meduc_ru_hs meduc_ru_mths smoked meduc_ru_unknown smoked_unknown if mwhite==1

Appendix C. Vital Statistics, Number and proportion per 1000, from year 1900 to year 1906

## VITAL STATISTICS.

No. 41.-DEATHS: Nummer and Proportion per 1,000, Calendar Years 1900 to 1906, AND AnNual. Average, 1901 to 1905, in the Registration Ahea, by Sgx AND AGe.
[From reports of the Burean of the Census, Department of Commerce and Labor.]
NUMBER OF DFATHR FROM ALL CAUSES

| Sex and age. | $\begin{aligned} & \text { Avernge, } \\ & \text { 1901-3, } \end{aligned}$ | 1900 | 1901 | 1902 | 1903 | 1904 | 1905 | 1906 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Sex: | 203,9042 | 205,909 | 276,020 | 273, \%25 |  |  |  |  |
| Fem | 215,668 | 253,910 | 242, 157 | 235,065 | 253,374 | 205, 102 | $\begin{aligned} & 202,912 \\ & 252,621 \end{aligned}$ | $299,819$ |
| inder 1 year | 100,268 | 111,687 | 97.477 | 98,575 | 95,857 | 102, 850 | 105, 583 | 123, 105 |
| 1 year | 22,325 | 25, 722 | 22,461 | 25,978 | 21,966 | 22,203 | 21,960 | 128,860 |
| 2 ye | 10,005 | 12, 129 | 10,031 | 10, 025 | 10,059 | 9, 209 | 9, 518 | 12, 188 |
| 3 years | 6,350 | 7,812 | 6,605 | 6, 475 | 6,429 | 6,323 | B, 916 | 7.430 |
| 4 years | 4,737 | 5,787 | 15,104 | 4,962 | 4,619 | 4,681 | 4,317 | 5,375 |
| Under 5 yea | 143, 681 | 164, 137 | 141,678 | 143, 515 | 139,940 | 145,902 | 147,384 | 186,978 |
| 5 to 9 years. | 13, 879 | 15,678 9 | 13, 932 | 13, 799 | 14,017 | 13, 7.4 | 12, 851 | 15, 317 |
| 10 to 14 year 15 | 8, 7143 | 9,144 14,498 | 8, 1136 | 8,163 13,709 | 8,733 | 15,368 | 8 8,835 | 10,463 |
| 10 to 19 years | 22,246 | 22,200 | 13,909 | 13, 209 | 14,511 | 15, 23,206 | 14,911 | 17,928 |
| 25 to 29 year | 24, 439 | 24, 573 | 24, 239 | 23, 512 | 24,639 | 25,336 | 24, 585 | 28, 633 |
| 30 to 34 ycars | 24, 169 | 23, 727 | 23,665 | 23, 382 | 25,063 | 25,237 | 24,506 | 28, 502 |
| 25 to 32 year | 25,352 | 23, 006 | 24, 456 | 24, 146 | 25,314 | 25, 49 | 26, 296 | 30, 790 |
| 40 to 44 years. | 24, 743 | 23,364 | 24,317 | 23, 797 | 24,672 | 25,787 | 25, 143 | 20, 101 |
| 45 to 49 year | 24, 008 | 27, 521 | 22, 802 | 22, 419 | 23,686 | 25, 487 | 25,948 | 30,703 |
| 50 to 54 year | 25, 706 | 24,203 | 24, 804 | 24,340 | 25,534 | 27, 152 | 26,671 | 31, 166 |
| 55 to 59 year | 25, 001 | 25,024 | 25, 308 | 24,654 | 26,030 | 27, 359 | 27,054 | 31, 989 |
| 60 to 64 year | 29,474 | 27,633 | 25, 491 | 27, 359 | 29,012 | $31,4 \mathrm{k}$ | 31,036 | 36, 109 |
| 65 to 69 years | 30, 382 | 29,123 | 29, 422 | 28, 427 | 30,335 | 31, 688 | 32, 037 | 38, 010 |
| 70 to 74 year | 30, 124 | 29,0es | 29, 161 | 2s, 196 | 29,736 | 22, 183 | 31,343 | 37,627 |
| 75 to 79 year | 25, 420 | 25,417 | 25,732 | 24, 474 | 26, 298 | 27,665 | 27,923 | 33, 501 |
| so to 84 yea | 19,446 | 18,843 | 19,494 | 18,147 | 19,222 | 20, 476 | 19,889 | 24,025 |
| 85 to 89 years. | 9,902 | 9,646 | 9,669 | 8,946 | 9,735 | 10,621 | 10, 841 | 13,071 |
| 90 to 94 year | 3, 118 | 3, 3113 | 3, 183 | 3,293 | 3,447 | 3,814 | 3, 601 | 4,179 |
| 96 years and | 1,118 | 1, 113 | 1,108 | 1,072 | 1,124 | 1,127 | 1,1588 | 1,393 |
|  | 1,801 | 1,927 | 2,252 | 1,909 | 2,000 | 1,743 | 1,013 | 1,805 |
| Aggregate | 529,650 | 239,939 | 518,207 | 505, 650 | 524, 415 | 851.354 | 545, 283 | 658, 105 |

PROPORTION PER 1,000 DEATHS.

| $536.2$ | $\begin{aligned} & 529.7 \\ & 470.3 \end{aligned}$ |
| :---: | :---: |
| 189.3 | 206. 8 |
| 42.2 | 49.5 |
| 18.9 | 22.5 |
| 12.0 | 14.5 |
| 8.9 | 10.7 |
| 271.3 | 304.0 |
| 25.8 | 29.0 |
| 16.4 | 16.9 |
| 27.4 | 26.9 |
| 42.0 | 41.2 |
| 46.1 | 45.5 |
| 45.6 | 43.9 |
| 47.8 | 45.6 |
| 46.7 | 43.3 |
| 45, 4 | 41.7 |
| 48.5 | 45.0 |
| 49.2 | 46.3 |
| 55.7 | 51.2 |
| 57.4 | 63.9 |
| 56.9 | 53.8 |
| 49.9 | 47.1 |
| 36.7 | 34.9 |
| 18.8 | 17.9 |
| 6.6 | 6.2 |
| 2.1 | 2.1 |
| 3.4 | 3.6 |
| 1,000.0 | 1,000.0 |




| 1 | $\begin{aligned} & 535.9 \\ & 46 \mathrm{~L} .1 \end{aligned}$ | 537.3 462.7 |
| :---: | :---: | :---: |
| 8 | 184.7 | 188. 6 |
| 2 | 41.9 | 40.4 |
| 7 | 19.2 | 17.7 |
| 7 | 12.3 | 11.5 |
| 8 | 8.8 | 8.5 |
| 2 | 206.8 | 261.6 |
| 1 | 26.8 | 25.0 |
| 0 | 16.7 | 17.0 |
| 0 | 27.7 | 28.1 |
| 1 | 42.4 | 42.1 |
| 3 | 47.0 | 46.0 |
| 0 | 45.9 | 45.8 |
| 5 | 48.3 | 48.0 |
| 8 | 47.0 | 46.8 |
| 1 | 45.2 | 46, 2 |
| 9 | 48.7 | 49.3 |
| 5 | 49.6 | 49.6 |
| 8 | 55.4 | 57.0 |
| 9 | 57.8 | 57.5 |
| 4 | 56.7 | 58.4 |
| 1 | 50.1 | 50.2 |
| 7 | 36.7 | 37.1 |
| 6 | 18.6 | 19.3 |
| 4 | 6.6 | 6.9 |
| 1 | 2.1 | 2.0 |
| 8 | 3.9 | 3.2 |
| 0 | 1,000,0 | 1,000.0 |




[^0]:    ${ }^{1}$ United Nations, Human Development Report 2005 (New York: UNDP, 2005), p. 4
    ${ }^{2}$ Definition of Infant mortality Rate, MedTerms.com, MedicineNet.com, http://www.medterms.com/script/main/art.asp?articlekey=3967
    ${ }^{3}$ Arthur Sullivan; Steven M. Sheffrin (2003). Economics: Principles in action. Upper Saddle River, New Jersey 07458: Pearson Prentice Hall. ISBN 0-13-063085-3
    ${ }^{4}$ Bill Hendrick (2009-11-04). "Preemies Raise U.S. Infant Mortality Rate". WebMD. http://www.webmd.com/baby/news/20091103/preemies-raise-us-infant-mortality-rate
    ${ }^{5}$ Gabriel Duc, "The crucial role of definition in perinatal epidemiology," Social and Preventive Medicine, Vol. 40, No. 6, November 1995

[^1]:    ${ }^{6}$ To simplify notation, further in paper African Americans identified as blacks and Caucasians identified as whites.
    ${ }^{7}$ Health factors well known medical conditions such as lower birthweight and shorter gestational period related to mother and child health. Finding or confirming these factors goes beyond of proposed paper, and remedy usually lies in healthcare advancements.
    ${ }^{8}$ Mario Sims, Tammy Sims, Marino Bruce, Journal of NMA, Apr 2007, 349-56 Urban Poverty and Infant Mortality Rate Disparities available at www.nmanet.org/images/uploads/Publications/OC349.pdf
    ${ }^{9}$ Almond, Hoynes, Schanzenbach NBER Working Paper No. 14306, Sept. 2008, Inside the War on Poverty: The Impact of Food Stamps on Birth Outcomes. Available at www.econ.ucdavis.edu/faculty/hoynes/working_papers/FSP_infants.pdf
    ${ }^{10}$ Gopal Singh, Michael Kogan, Pediatrics Vol. 119 No. 4 April 2007, pp. e928-e939 Persistent Socioeconomic Disparities in Infant, Neonatal, and Postneonatal Mortality Rates in the United States, 1969-2001 Available at http://pediatrics.aappublications.org/cgi/content/full/119/4/e928

[^2]:    ${ }^{11}$ Farley, Mason, Rice, Habel Pediatric and Perinatal Epidemiology 2006 May; 20(3):188-200 The relationship between the neighborhood environment and adverse birth outcomes
    ${ }^{12}$ Sohler, Arno, Chang, Fang and Clyde Schechter Journal of Urban Health Volume 80, Number 4, Dec 2003, 650657 Income inequality and infant mortality in New York City
    ${ }^{13}$ Andrew Leigh and Christopher Jencks, Journal of Health Economics 26 (2007) 1-24 Inequality and mortality: Long-run evidence from a panel of countries. http://people.anu.edu.au/andrew.leigh/pdf/InequalityMortality.pdf ${ }^{14}$ Mario Sims and Yolanda Rainge, Journal of National Med. Assoc. 2002 June; 94(6): 472-479, Urban poverty and infant-health disparities among African Americans and whites in Milwaukee
    ${ }^{15}$ Marian F. MacDorman, Ph.D., and T.J. Mathews, M.S., CDC, "Behind International Rankings of Infant Mortality: How the United States Compares with Europe". http://www.cdc.gov/nchs/data/databriefs/db23.htm

[^3]:    ${ }^{16}$ StataCorp LP, www.stata.com, In the paper I use Stata version 10.1
    ${ }^{17}$ CDC NCHS, Division of Vital Statistics, 2004 Birth Cohort Linked Birth/Infant, Death Data Set, Dataset and Guide Available at http://www.cdc.gov/nchs/data_access/Vitalstatsonline.htm
    ${ }^{18}$ Non-residents were excluded or 65 occurrences
    ${ }^{19}$ U.S. Census Bureau, County and City Data Book http://www.census.gov/statab/ccdb/ccdbstcounty.html

[^4]:    ${ }^{21}$ I picked State of California as the base state for my model. I believe California's large diverse population makes it natural choice for national level model.

[^5]:    ${ }^{22}$ A Friede et al. Public Health Report, Mar-Apr 1987, Young maternal age and infant mortality: the role of low birth weight. Available at http://www.ncbi.nlm.nih.gov/pmc/articles/PMC1477817/
    ${ }^{23}$ Author finds the any policy, which restricts the maternal age or the number of kids, violates very core of human rights (even if such policy is effective in reducing IMR it creates enormous drawback form the moral standpoint). One could argue that permitting mother to have child outside of favorable maternal age creates even higher suffering for mother if her child dies. The only very naive solution I could think of is to compare decrease in utility of not having a baby outside of maternal age to decrease in utility of the death of child multiplied by event probability in the population.
    ${ }^{24}$ Mothers from age 26 to 30 were chosen as a base group.
    ${ }^{25}$ In the dataset $3.365 \%$ of black mothers were between ages 11 and 16 , compare to only $1.355 \%$ of their white counterparts or 2.5 times higher in the age bracket where IMR averages between 20 and 10 per 1000

[^6]:    ${ }^{26}$ Estimate has very high $R^{2}=0.6754$, which is not surprising because it relies on IMR data aggregated by age

[^7]:    ${ }^{27}$ Education level 'less than high school' was chosen as base for this model.

[^8]:    ${ }^{28}$ Rebecca Din-Dzietham, Irva Hertz-Picciotto, Infant mortality differences between whites and African Americans: the effect of maternal education. Am J Public Health. 1998 April; 88(4): 651-656. Available at http://www.ncbi.nlm.nih.gov/pmc/articles/PMC1508444/pdf/amjph00016-0115.pdf
    ${ }^{29}$ Office of Minority Health, African American Profile, Available at http://minorityhealth.hhs.gov/templates/browse.aspx?lv|=3\&|vlid=23

[^9]:    ${ }^{30}$ One might speculate of possible reasons for such behavior. For example: reducing medical expenses, lack of time for regular prenatal visits during the day, or bias towards healthcare professionals. No matter what reason is woman believes that benefits of not having prenatal visits outweigh the risks.
    ${ }^{31}$ From the dataset it's hard to say whether self-assessment of white woman, on average, is better than black counterpart, or it's just imprudence. Probably both factors play the role, considering that blacks give birth at younger age and thus more prone to error in judgment.

[^10]:    ${ }^{32}$ Sadly, I couldn't get county statistics on obstetricians and gynecologists (OB/GYN). I waited for copy of American Medical Association annual edition of Physician Characteristics and Distribution, only to find out that edition doesn't list OB/GYN professions separately from other fields at the county level.
    ${ }^{33}$ Federal Information Processing Standards, read more at http://www.itl.nist.gov/fipspubs/

[^11]:    ${ }^{35}$ James Cramer, Racial and ethnic differences in birthweight: the role of income and financial assistance, Demography. 1995 May; 32(2):231-47.
    ${ }^{36}$ Based on data $10.9 \%$ of white mothers smoked during the pregnancy compared to only $8.2 \%$ for black

[^12]:    ${ }^{37}$ Mario Sims et al, Journal Of The National Medical Association, Vol. 99, No. 4, April 2007, Urban Poverty and Infant Mortality Rate Disparities, http://www.nmanet.org/images/uploads/Publications/OC349.pdf
    ${ }^{38}$ Aber et al, Annual Review Public Health. 1997. 18:463-83, The Effects Of Poverty
    On Child Health And Development http://homepages.nyu.edu/~dc66/pdf/res_Annual_Review_aber_etal.pdf
    ${ }^{39}$ To be conclusive, such experiment need substantial funding, and can resemble a study conducted by UK government in late 1970s and known as Black Report. Available at http://www.sochealth.co.uk/Black/black.htm

[^13]:    ${ }^{40}$ Of course there's possibility that both don't have a job but this possibility lower compare to single woman. We can also take into account that men make more on average than women
    ${ }^{41}$ Here I mean not a financial, but moral support and help during the pregnancy. We usually call it: taking care of pregnant wife.
    ${ }^{42}$ In 2004-69.5\% of white mothers reported of being married compared to only $31.2 \%$ of black counterparts

[^14]:    ${ }^{44}$ Based on my estimates records for blacks more often have missing data for all three factors.
    ${ }^{45}$ Omitting unknown variables from the model increase the gap between blacks and whites to 6.45 per 1000 . Removing records with missing data from the sample reduces the gap to 5.88 per 1000. Thus missing records increase the gap between black and whites and there's more missing data on blacks compare to whites. Finding a solution to missing data can further extend our understanding of the adverse effect. There are several modern statistical methods help to fill-in missing data. Unfortunately none of them as good as actual data.
    ${ }^{46}$ According to my calculations in $200466 \%$ of infants' deaths occurred in neonatal period, and $40 \%$ of deaths occurred in the first day of life.
    ${ }^{47}$ For comparison neonatal mortality is $68.80 \%$ for Asians and $68.63 \%$ for Hispanics

[^15]:    ${ }^{48}$ Generally speaking there is more than just income we could include individual ability, family background, level can be affected by different rate of access to funds, but we'll focus on income as one of socioeconomic factors. ${ }^{49}$ Black mothers the only one experiencing higher estimates with increased level of education in neonatal period. Hispanics and Asians have estimates similar to whites.
    ${ }^{50}$ Growth in IMR with trimester prenatal care started in postneonatal period perhaps can be explained by how much resources were devoted to an infant. Newborns with severe health issues died during neonatal period. Rest of infants depended on resources and efforts devoted by their mothers or parents. Starting prenatal care yearly can be viewed as resource proxy. Dedicated mothers and mothers with greater resource available likely to start prenatal care earlier and keep putting more resources and efforts after child was born.

[^16]:    ${ }^{51}$ For more information please see Appendix C. Copy of Vital Statistics Report No 41, Bureau of Census, Department of Labor and Commerce, Number and Proportion of Deaths for Period Between 1900-1906.
    ${ }^{52}$ CDC MMW Report October 01, 1999 / 48(38);849-858 Achievements in Public Health, 1900-1999: Healthier Mothers and Babies, available at http://www.cdc.gov/mmwr/preview/mmwrhtml/mm4838a2.htm
    ${ }^{53}$ We could estimate socioeconomic factors influencing gap more precisely if we could use direct measures of income and utilization of health care (including insurance coverage). Unfortunately on individual level CDS data gives us only proxies for such measures and county level data is insignificant due to many limitations described in the paper

